



Inverse problems in functional brain imaging

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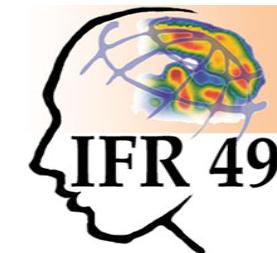
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1: CEA/NeuroSpin/LNAO



2: IFR49



Outline

- Introduction to functional brain imaging
 - Magneto/Electro-encephalography
 - functional MRI
- Mapping brain activity
 - GLM framework (Statistical Parametric Mapping)
 - Spatial regularization [VB optimization]
- Probe brain dynamics in fMRI
 - FIR modeling
 - Temporal regularization [EM/SAEM algorithms]
- Joint detection-estimation
 - Unify both questions
 - Spatio-temporal regularization [MCMC methods]



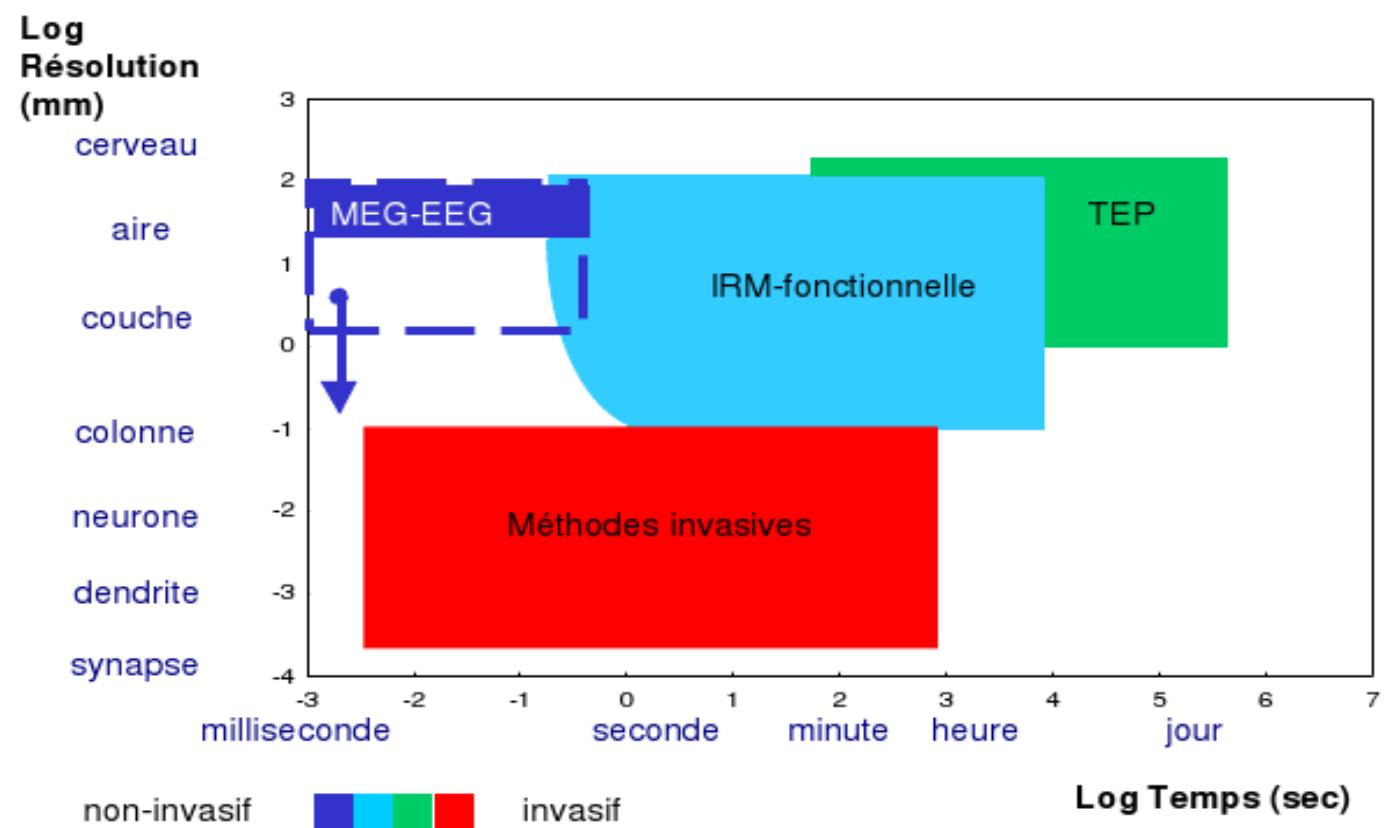
Functional brain imaging

Imagerie Fonctionnelle Cérébrale

- Etude du cerveau en action
- Nombreuses applications cliniques et en Sciences Cognitives
- À l'interface entre les Sciences de la Vie et les Sciences pour l'Ingénieur

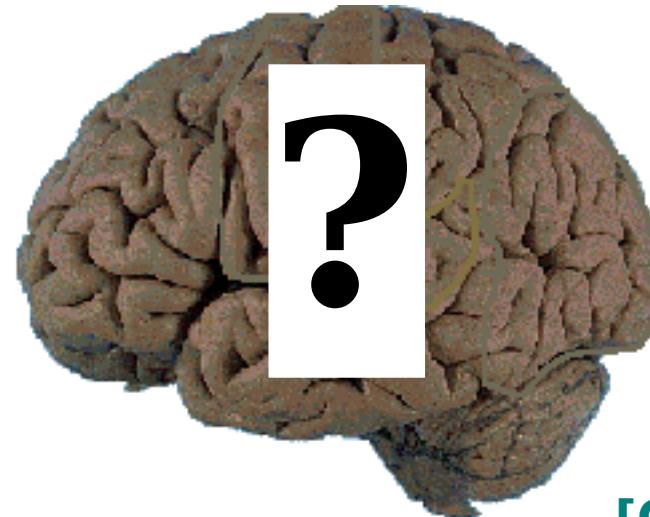
Techniques

Où et Quand ?

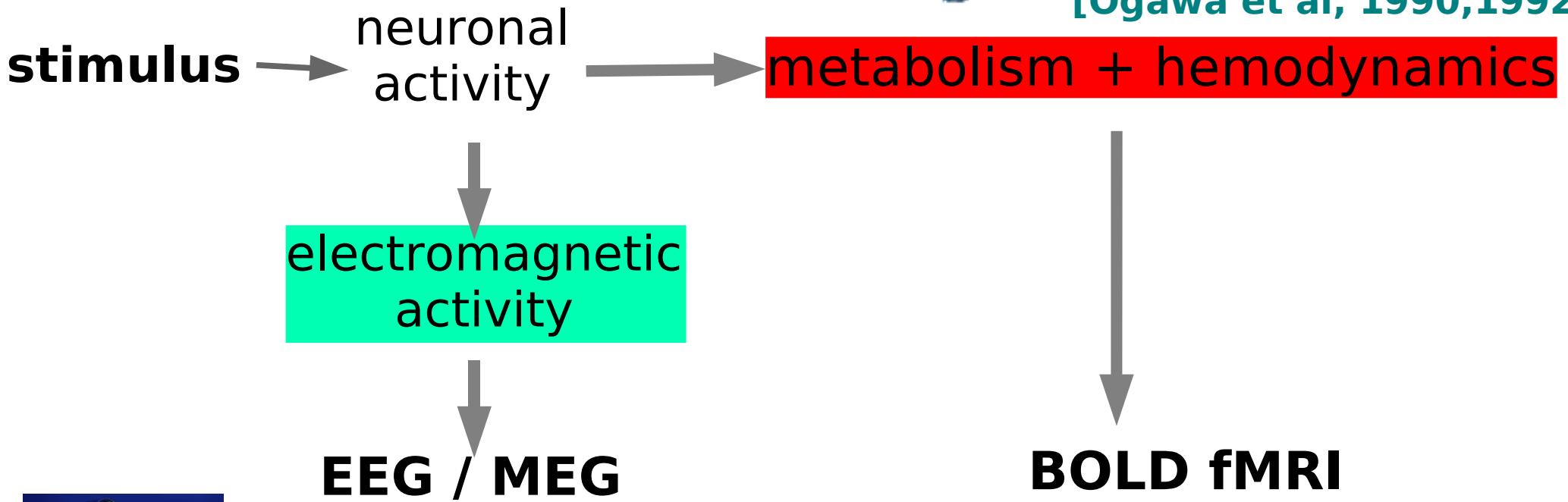


Functional brain imaging

Probe brain dynamics
non-invasively



[Ogawa et al, 1990, 1992]



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MEG/EEG

Électroencéphalographie (EEG)

Activité électrique neuronale

Résolution temporelle : $\sim 1\text{ms}$

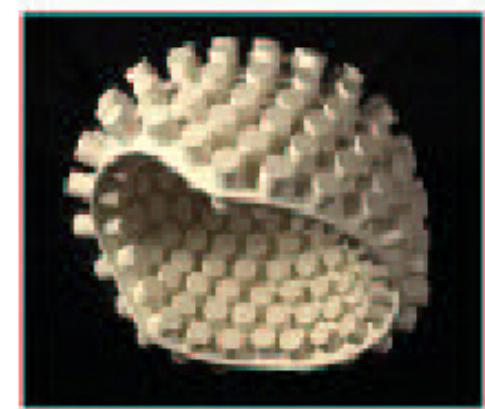
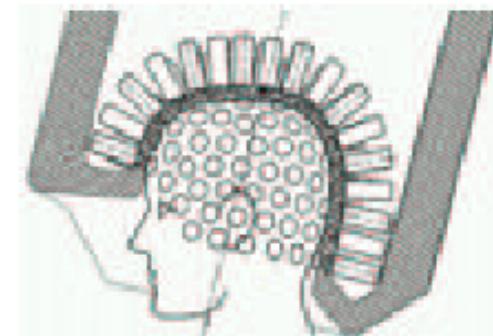


EEG : mesure du potentiel électrique

Ordre de grandeur : qq $\mu\text{-volts}$

Capteurs : électrodes

Magnétoencéphalographie (MEG)



MEG : mesure du champ magnétique

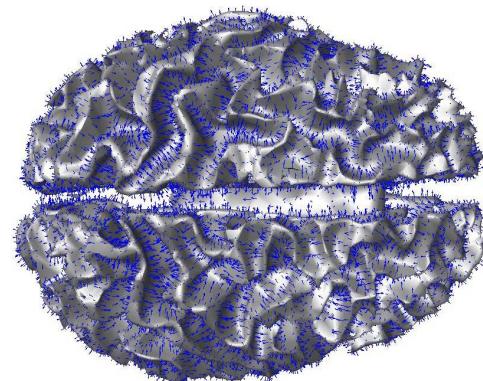
Ordre de grandeur : 10^{-13} Tesla

Capteurs SQUID couplés à des bobines



Posthumous honor to Line Garner

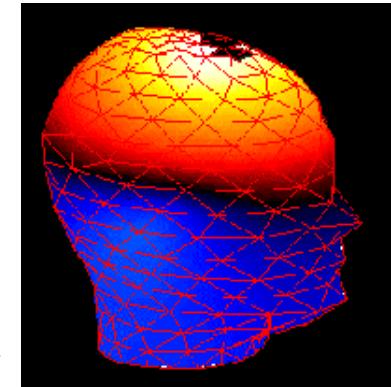
MEG/EEG source reconstruction



Distributed
Source model

J

Inverse
procedure



Data

K

Forward
modelling

$$Y = K * J + E$$

$$Y = KJ + E$$

[nxt] [nxpIpxt] [nxt]

n : number of sensors
p : number of dipoles
t : number of time samples

- under-determined system
- priors required

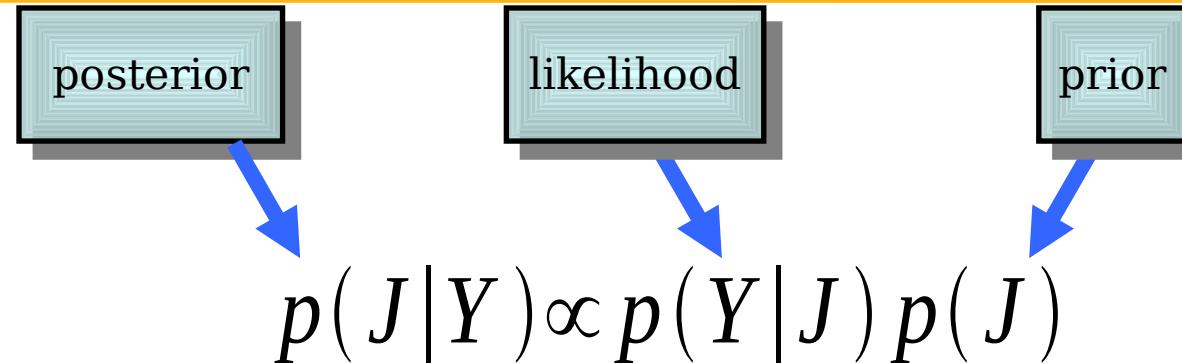


Bayesian
framework

[Mattout et al., Neuroimage, 2006]



MEG/EEG source reconstruction

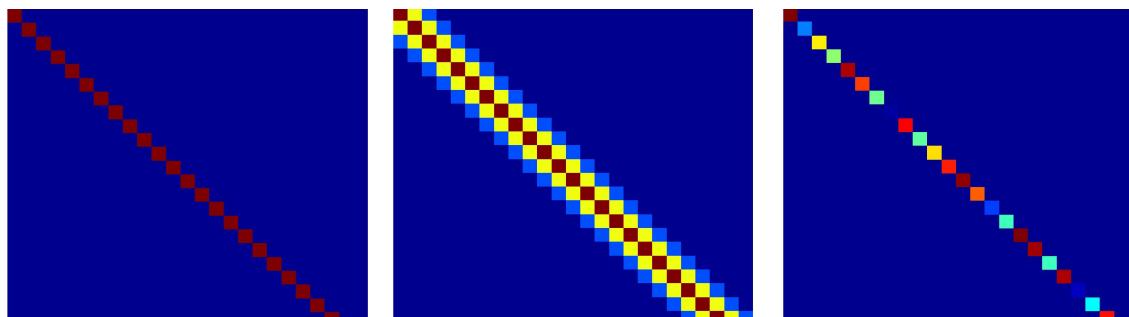


$$U_{MAP}(J) = \|C_{e^{-1/2}}(Y - KJ)\|^2 + \lambda \|WJ\|^2$$

likelihood WMN prior

$$p(J) \sim N(0, C_j) \quad C_j^{-1} = \lambda W^T W$$

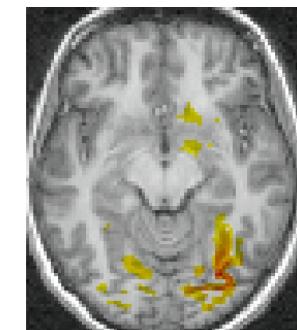
2-level hierarchical model:



minimum norm smoothness prior functional prior

$$Y = KJ + E_1 \quad E_1 \sim N(0, C_e)$$

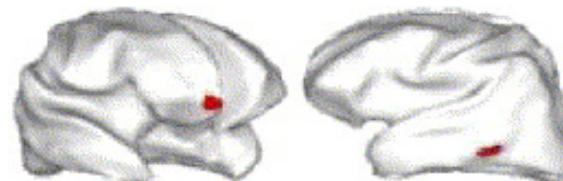
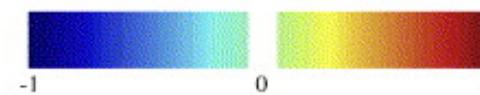
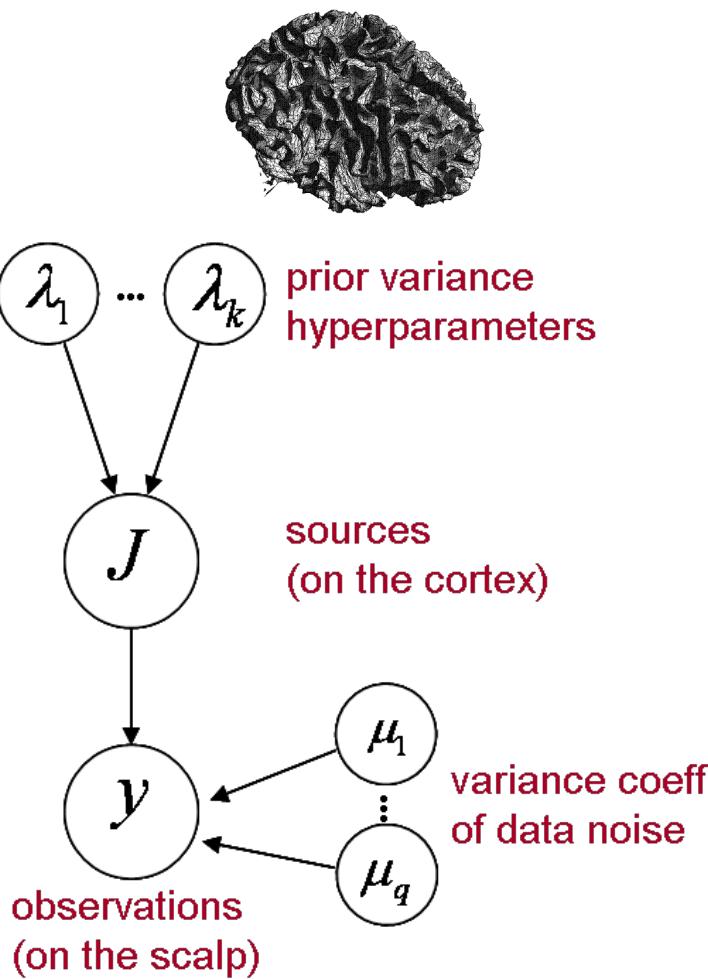
$$J = 0 + E_2 \quad E_2 \sim N(0, C_p)$$



[Mattout et al., Neuroimage, 2006]



MEG/EEG source reconstruction



(a) Source locations



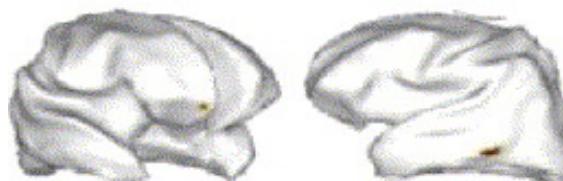
(b) Invalid prior location



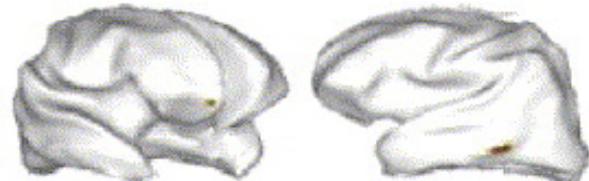
(c) WMN solution under the smoothness prior



(d) ReML solution under the smoothness prior



(e) ReML solution under the smoothness and valid priors



(f) ReML solution under the smoothness, valid and invalid priors

[Mattout et al., Neuroimage, 2006]



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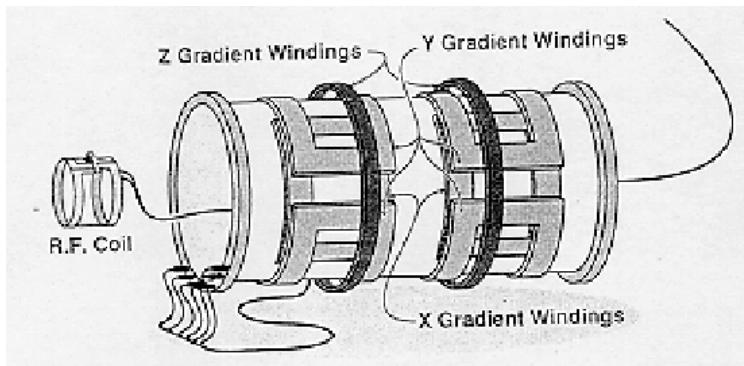


Magnetic Resonance Imaging

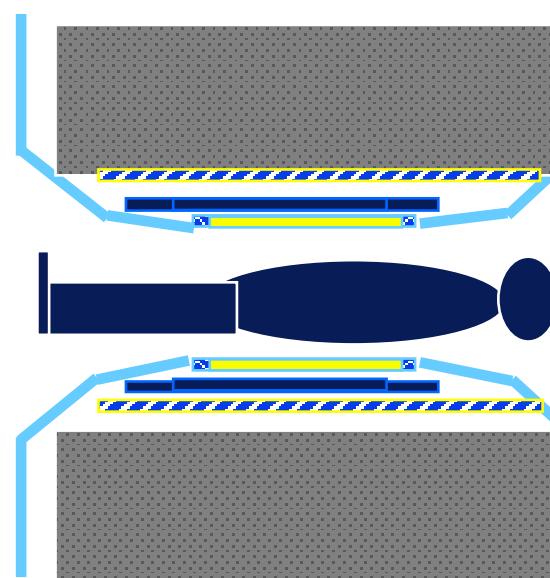


⇒ High magnetic field

3T



⇒ Auxiliary coils:
the « gradients »



⇒ Émetteur/récepteur RF



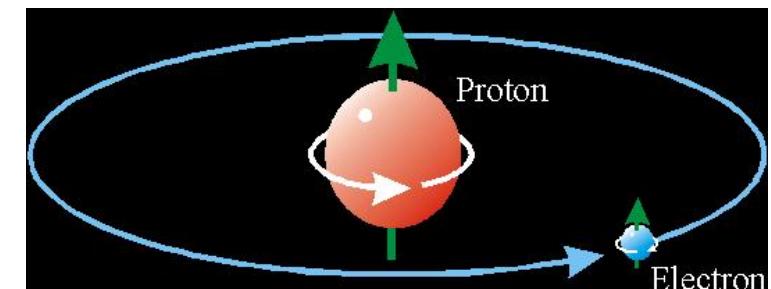
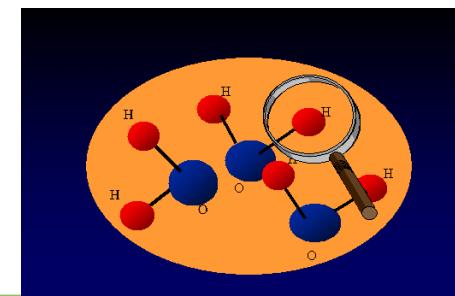
Atomes en RMN

Éléments	Abondance biologique
Hydrogen (H)	0.63
Sodium (Na)	0.00041
Phosphorus (P)	0.0024
Carbon (C)	0.094
Oxygen (O)	0.26
Calcium (Ca)	0.0022
Nitrogen (N)	0.015

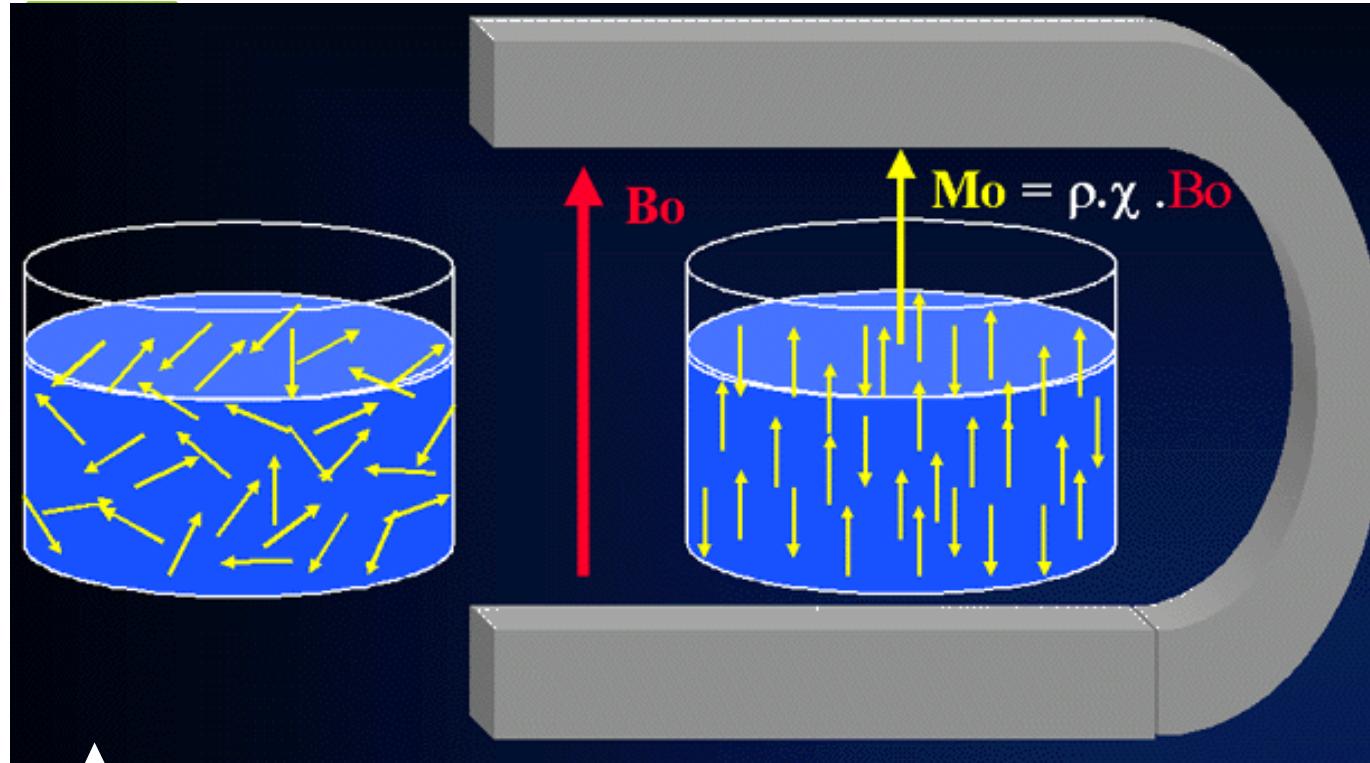


Éléments Utilisé en RMN	Symbolé	Abondance dans le corps humain
Hydrogen	^1H	99.985
	^2H	0.015
Carbon	^{13}C	1.11
Nitrogen	^{14}N	99.63
	^{15}N	0.37
Sodium	^{23}Na	100
Phosphorus	^{31}P	100
Potassium	^{39}K	93.1
Calcium	^{43}Ca	0.145

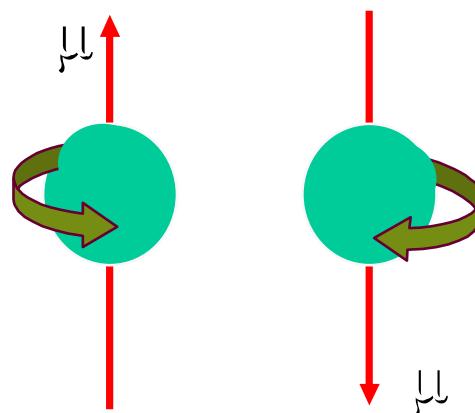
Choix des **protons des atomes d'hydrogène** de l'eau pour l'imagerie



Choix de l'aimant statique



Alignment des moments magnétiques suivant 2 directions:



- direction parallèle: orientation dans la direction de B_0
- direction antiparallèle: orientation dans la direction opposée à B_0

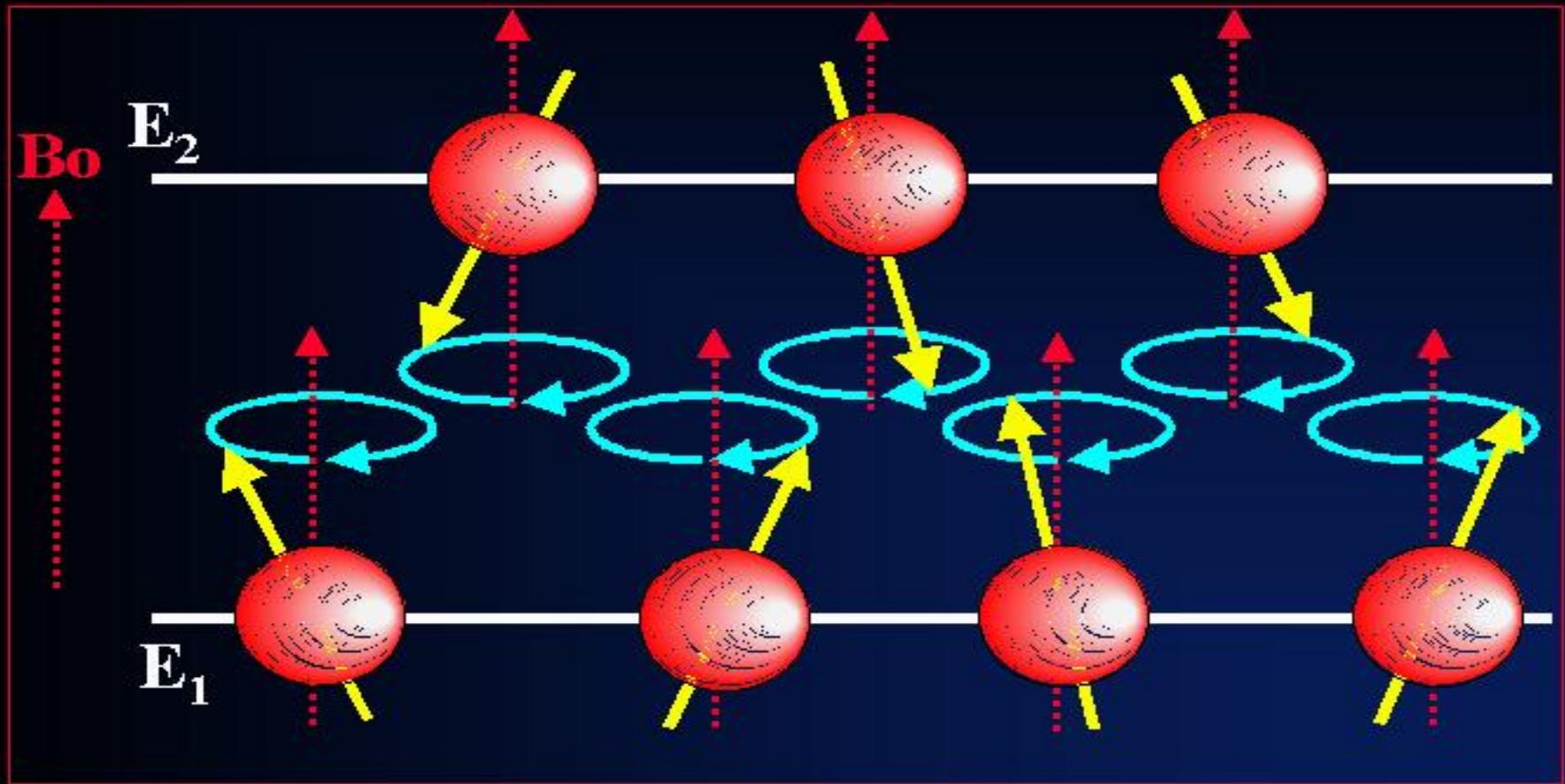
Mouvement de précession

$$\omega = \gamma B$$

Pour les protons 42.58 MHz/T

Spins des protons

Les deux états d'énergie magnétique du noyau d'Hydrogène

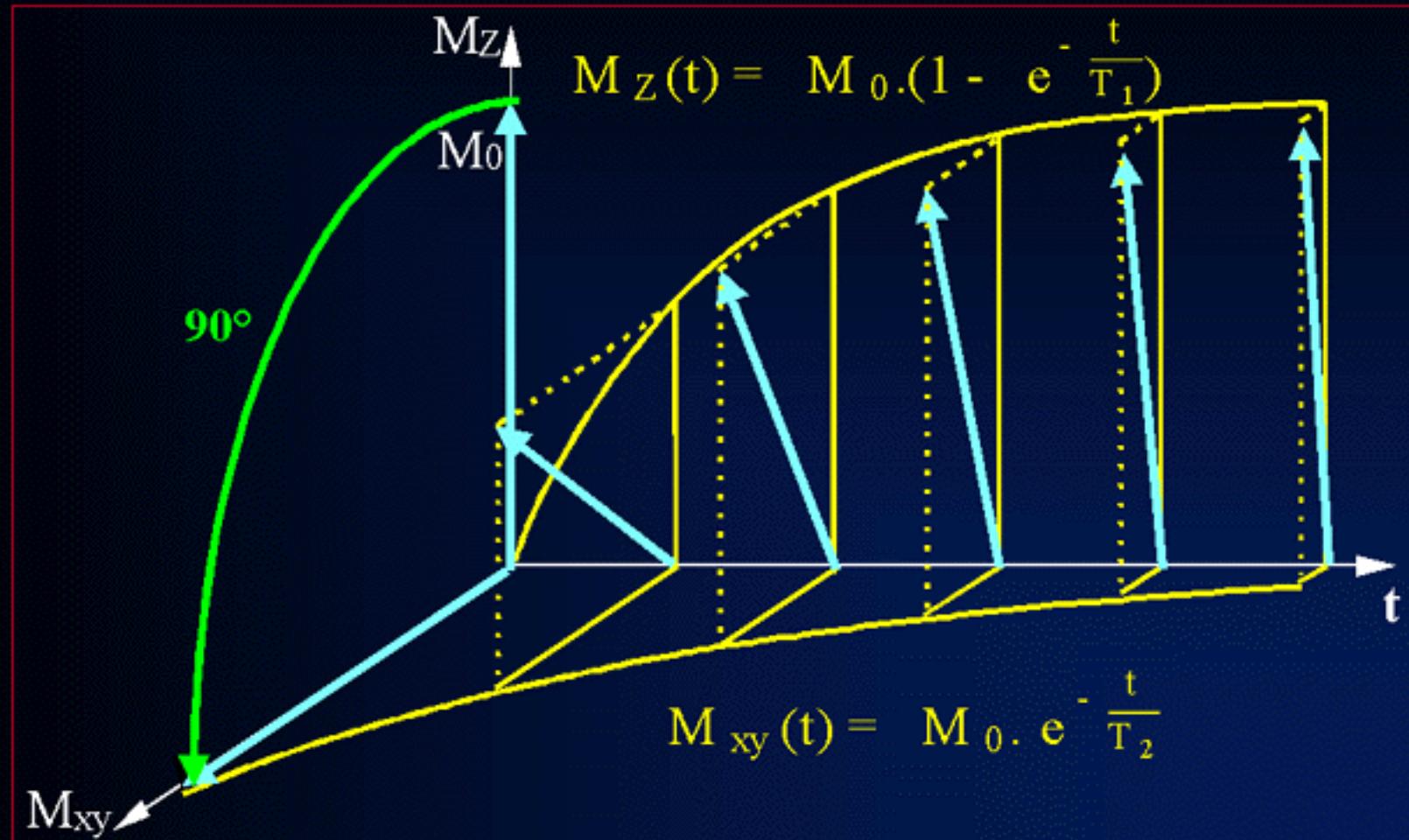


en présence d'un champ magnétique B_0



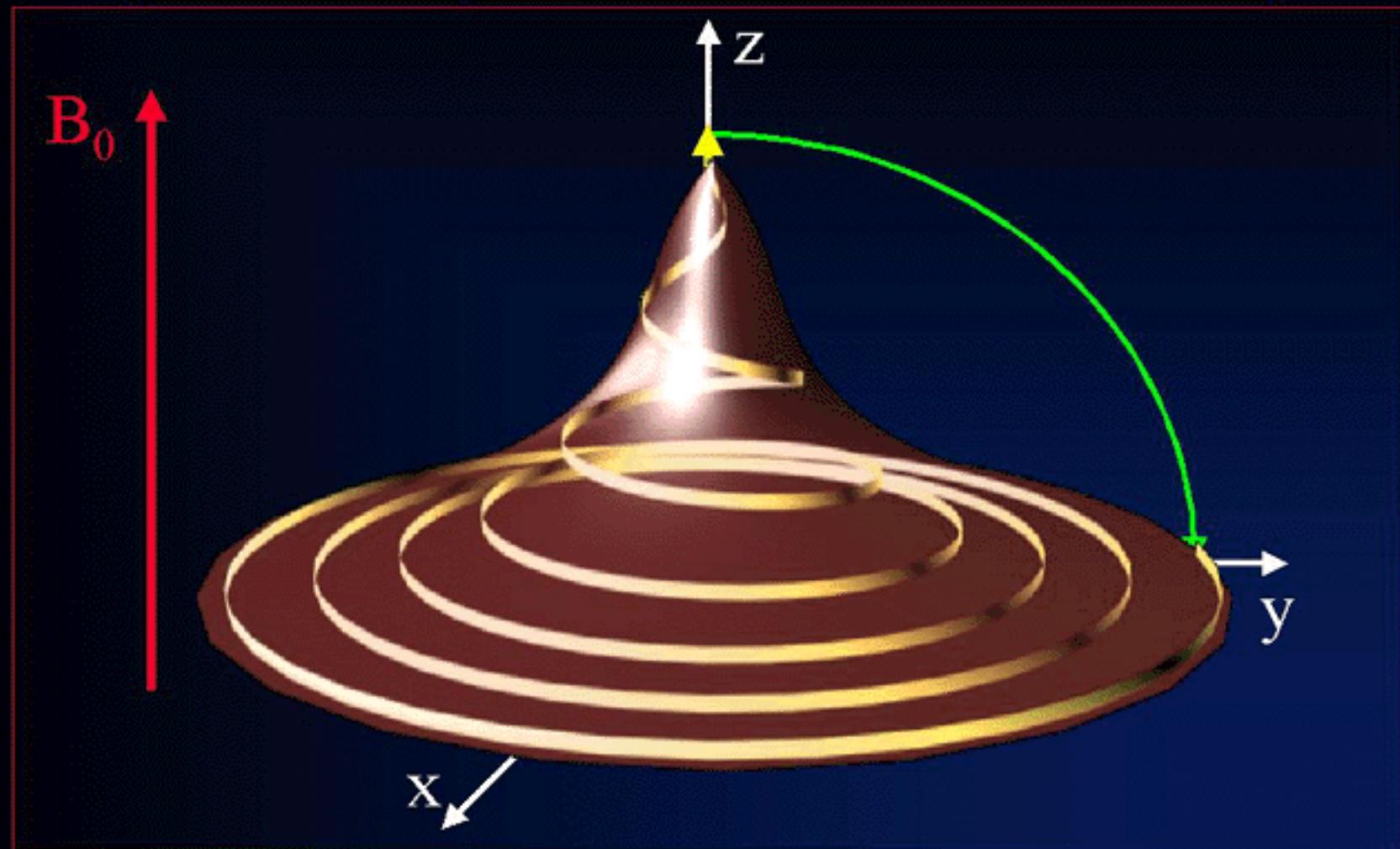
Phénomène de RMN

Relaxation



Phénomène de RMN

Retour à l'Equilibre de l'Aimantation Nucléaire

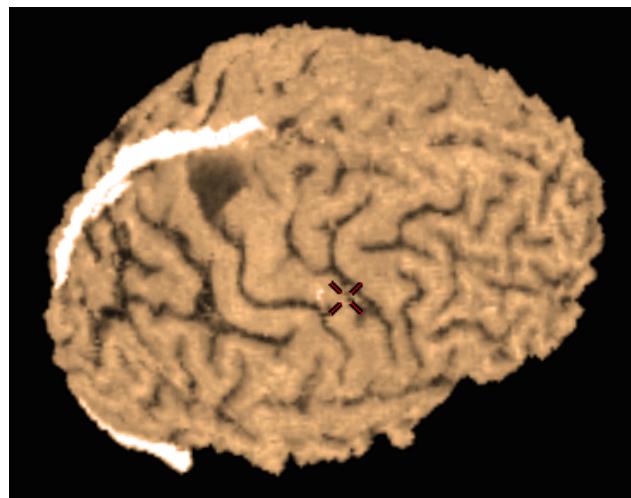
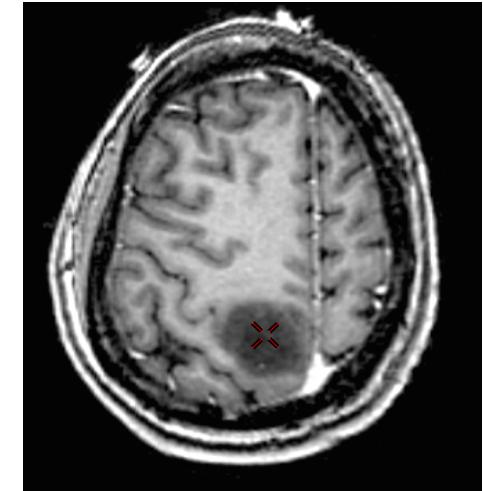
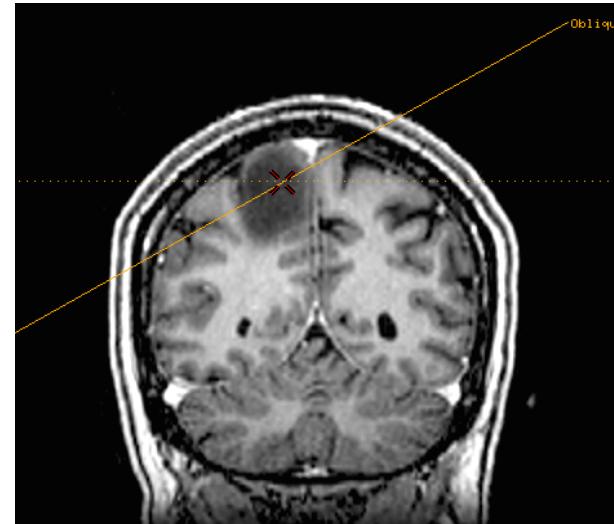
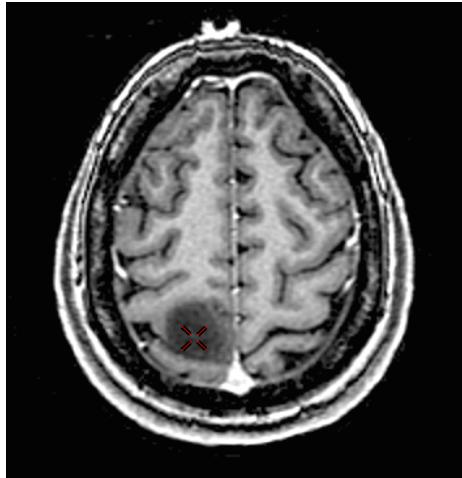


Temps de relaxation

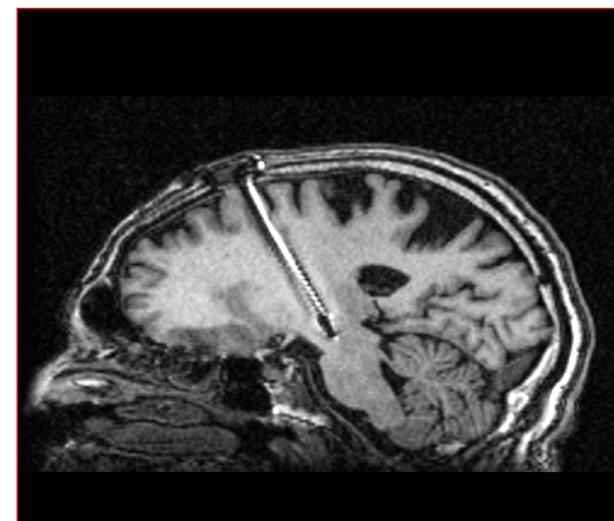
Ordres de Grandeurs des Temps de Relaxation à 1,5 T

Temps de Relaxation	T1	T2
Tissus Humains		
Liquide Céphalo-rachidien	2500 ms	2000 ms
Substance Grise	900 ms	90 ms
Substance Blanche	750 ms	80 ms
Graisse	300 ms	40 ms

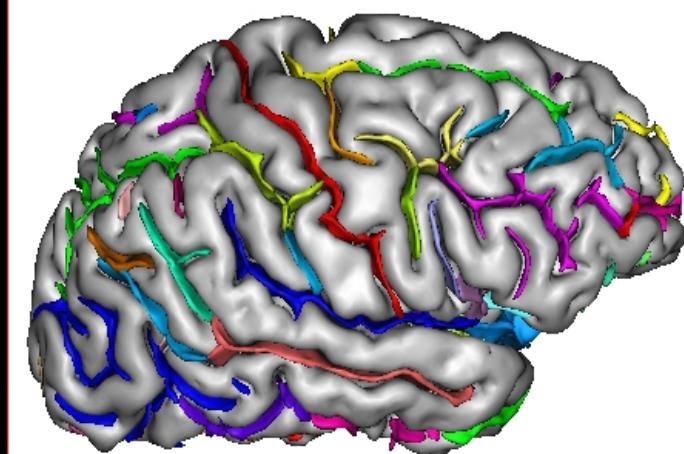
Anatomical MRI



Brain tumor



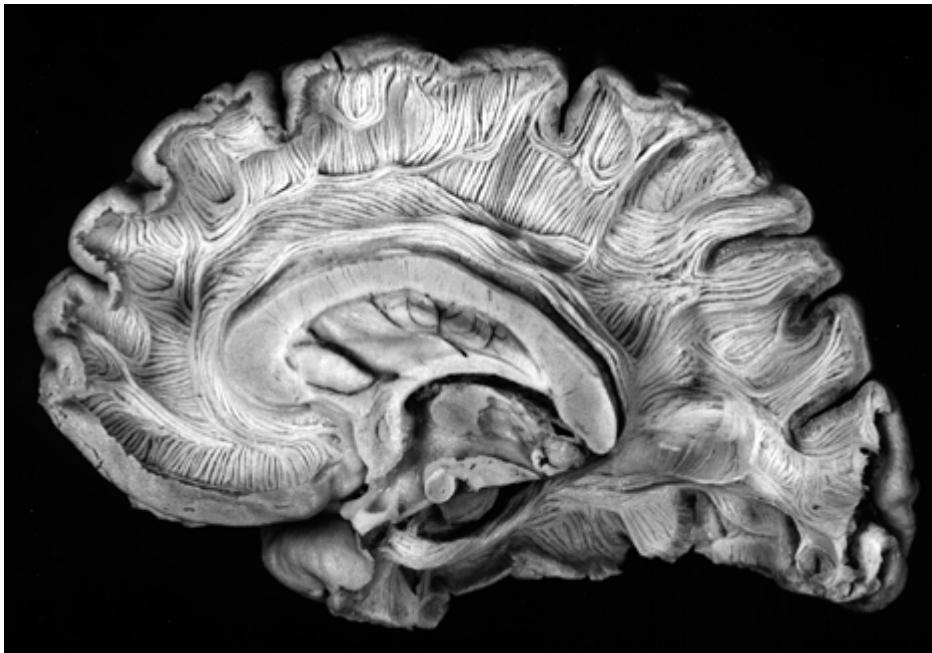
Post-op checking



Sulci recognition



diffusion MRI



[Virtual Hospital, 1998]



[Elkouby, 2005]

- ✗ **in vivo imaging of the water brownian motion**
- ✗ **Get access to the structural connectivity**

Mouvement brownien
[Einstein, 1905]

Spectro diffusion
[Stejskal-Tanner, 1965]

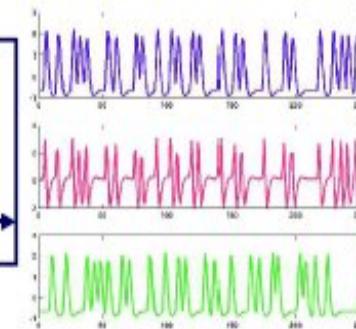
IRM de diffusion
[Le Bihan, 1986]



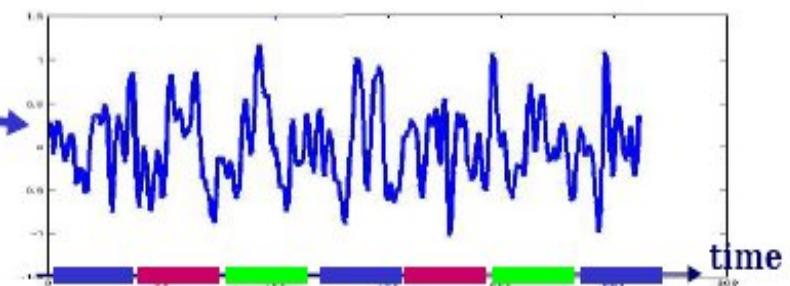
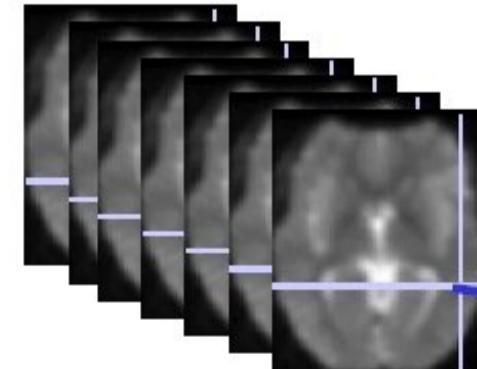
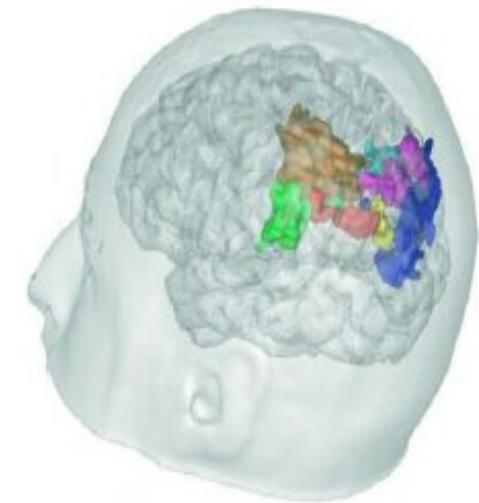
Functional MRI



Experimental Paradigm



Neuronal activation



Measured 4D signal :

BOLD = Blood Oxygenation Level Dependent signal

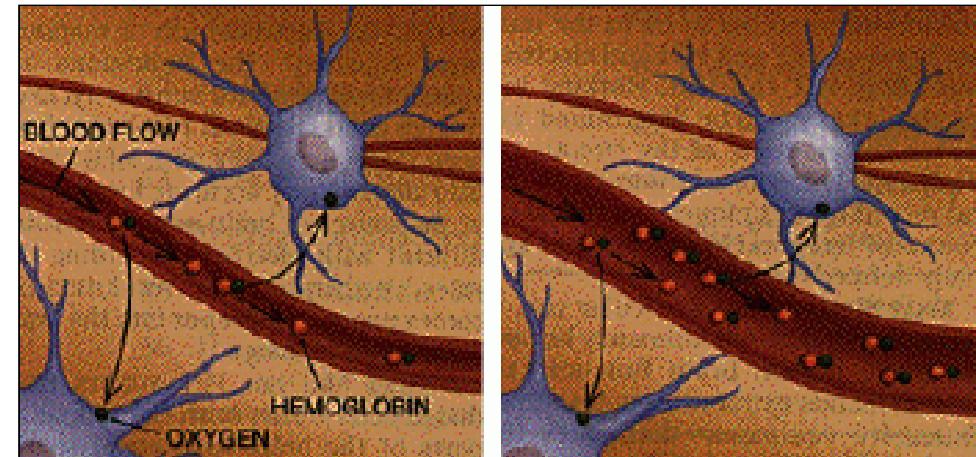


Le signal BOLD

[Ogawa et al, 1990,1992]

Produit de contraste intrinsèque :
 Oxyhémoglobine (HbO_2) : diamagnétique
 Désoxyhémoglobine (Hb) : paramagnétique

Détectable en IRMf

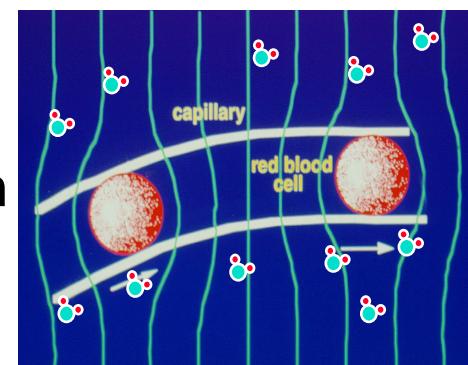


Activation cérébrale



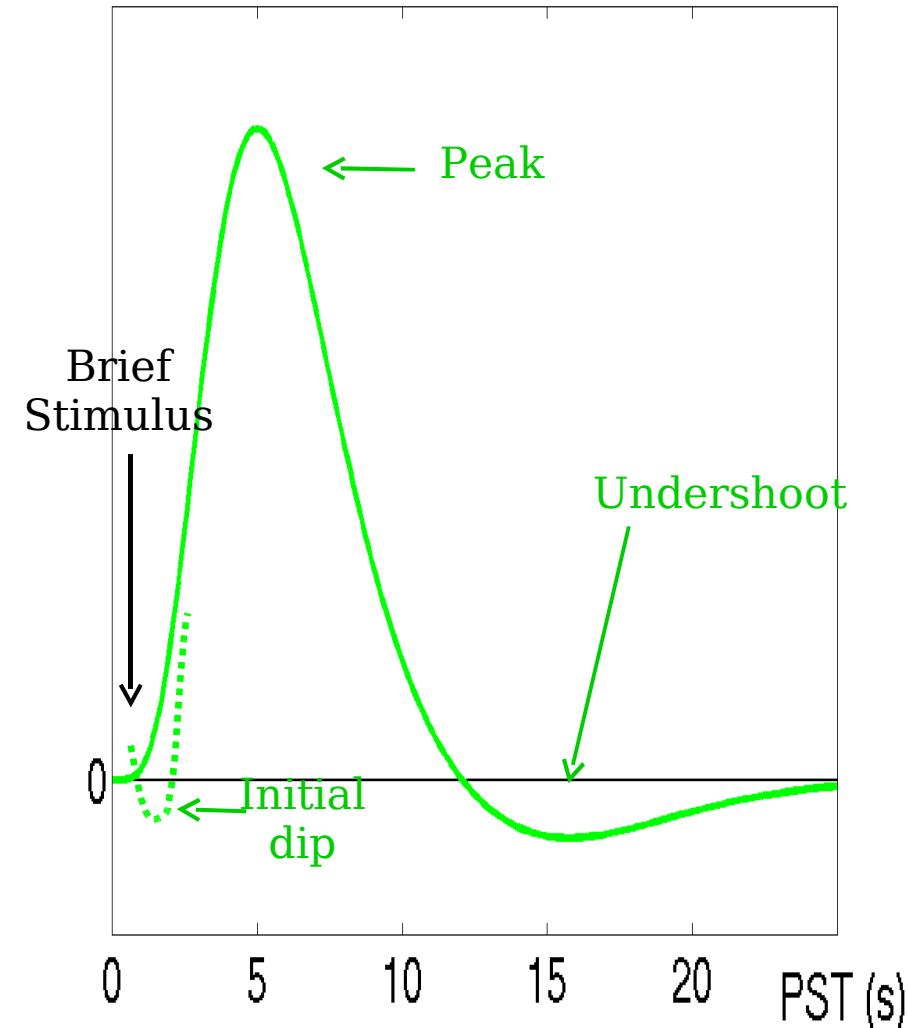
Légère augmentation de la consommation O_2 , accompagnée d'un fort afflux de sang oxygéné

Conséquence : augmentation de la concentration en sang oxygéné (HbO_2) des vaisseaux proches des neurones actifs



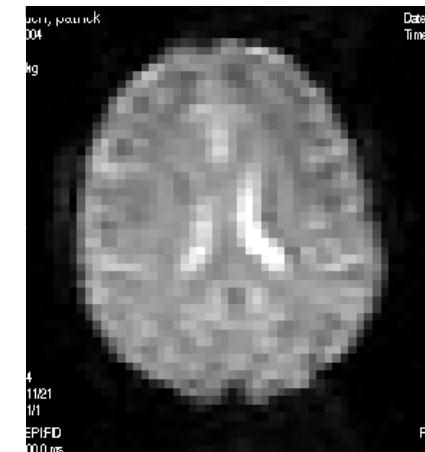
Hemodynamic Response Function

- Function of blood oxygenation, flow, volume [Buxton et al. 98]
- Peak (max. oxygenation) 4-6s poststimulus;
- Baseline after 20-30s
- Initial undershoot can be observed [Malonek & Grinvald. 96]
- ... but differences across: other regions [Schacter et al. 97]
- individuals [Aguirre et al. 98]

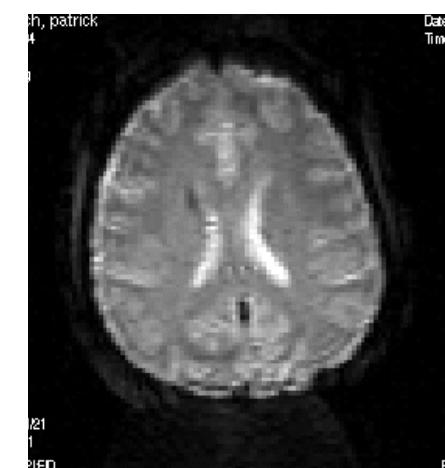


The truth about fMRI data

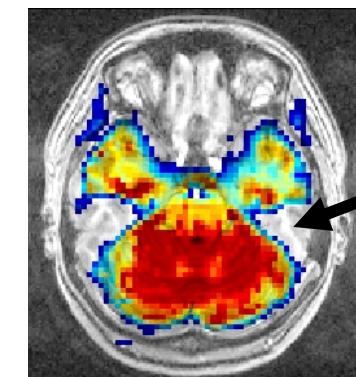
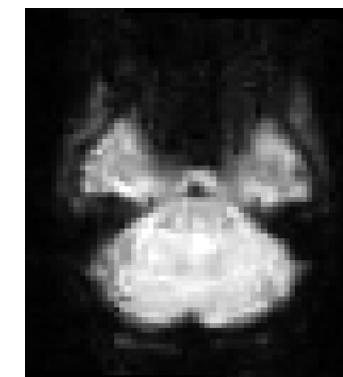
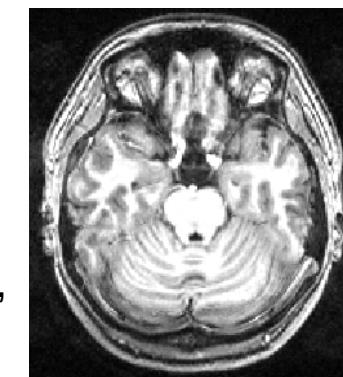
64x64 Pixels ~ 3 x 3 mm



128x128 Pixels ~ 1.5 x 1.5mm



I. They are distorted



II. They are noisy

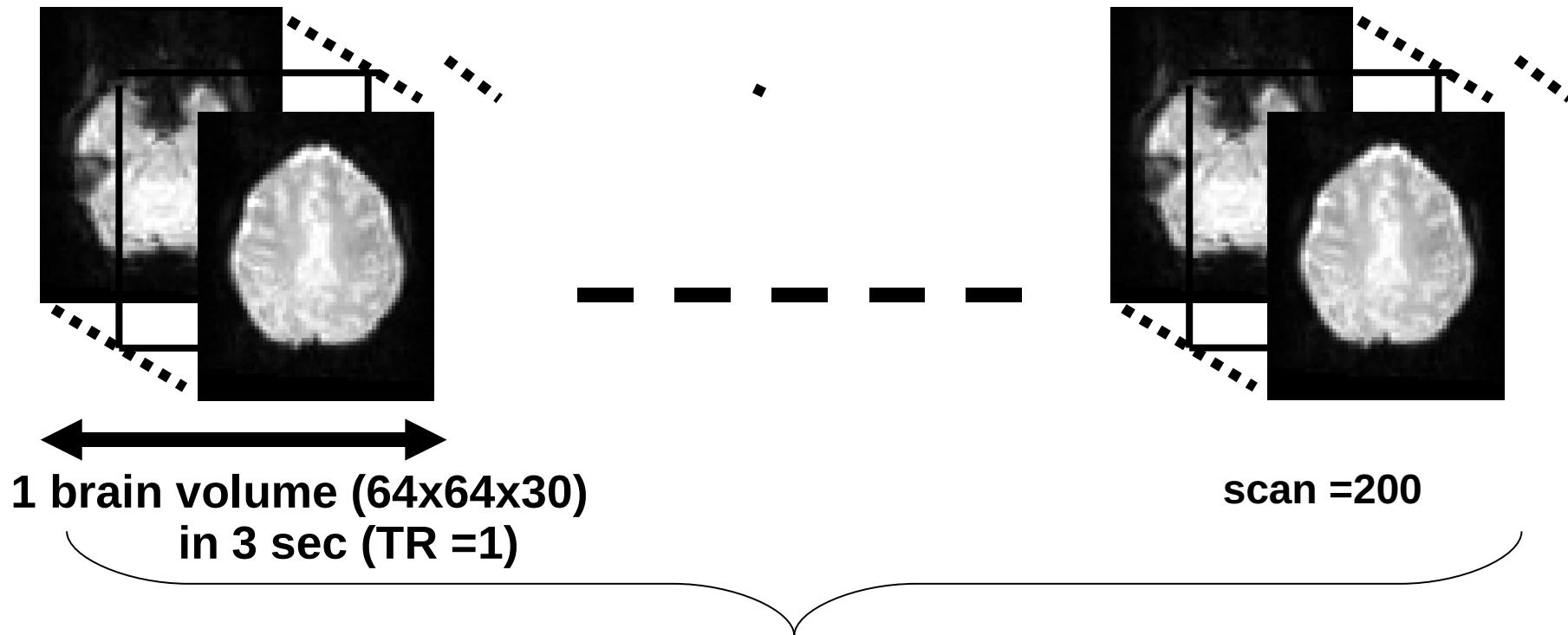
III. They don't have signal everywhere in the brain

IV. They depend on many parameters: T2*, B0, TE, ...



The truth about fMRI data

V. They are big ... and are getting bigger



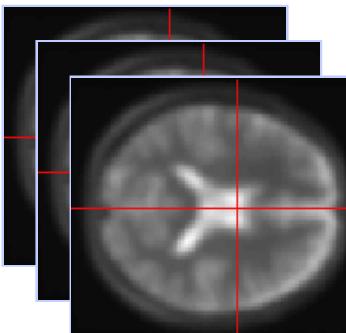
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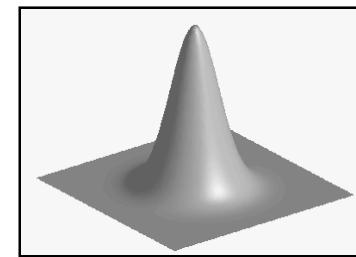


fMRI processing pipeline

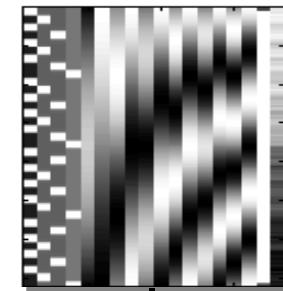
Image time-series



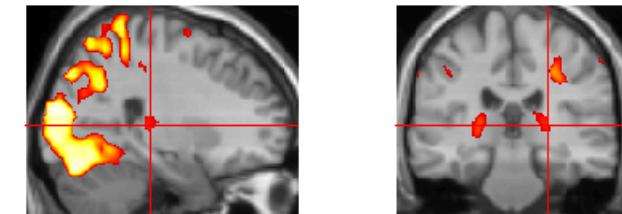
Spatial filter



Design matrix



Statistical Parametric Map

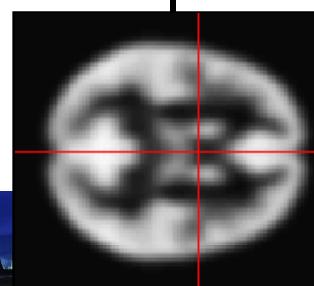


Realignment

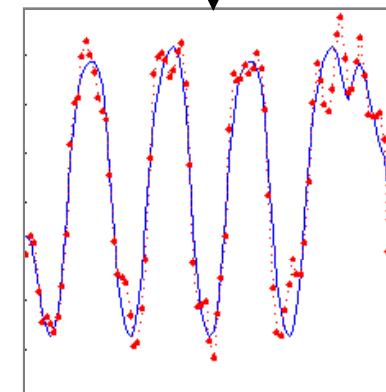
Smoothing

General Linear Model

Normalisation

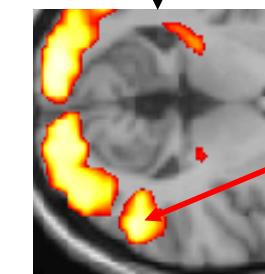


Anatomical
reference



Parameter estimates

Statistical
Inference



RFT

$p < 0.05$

fMRI example

cea

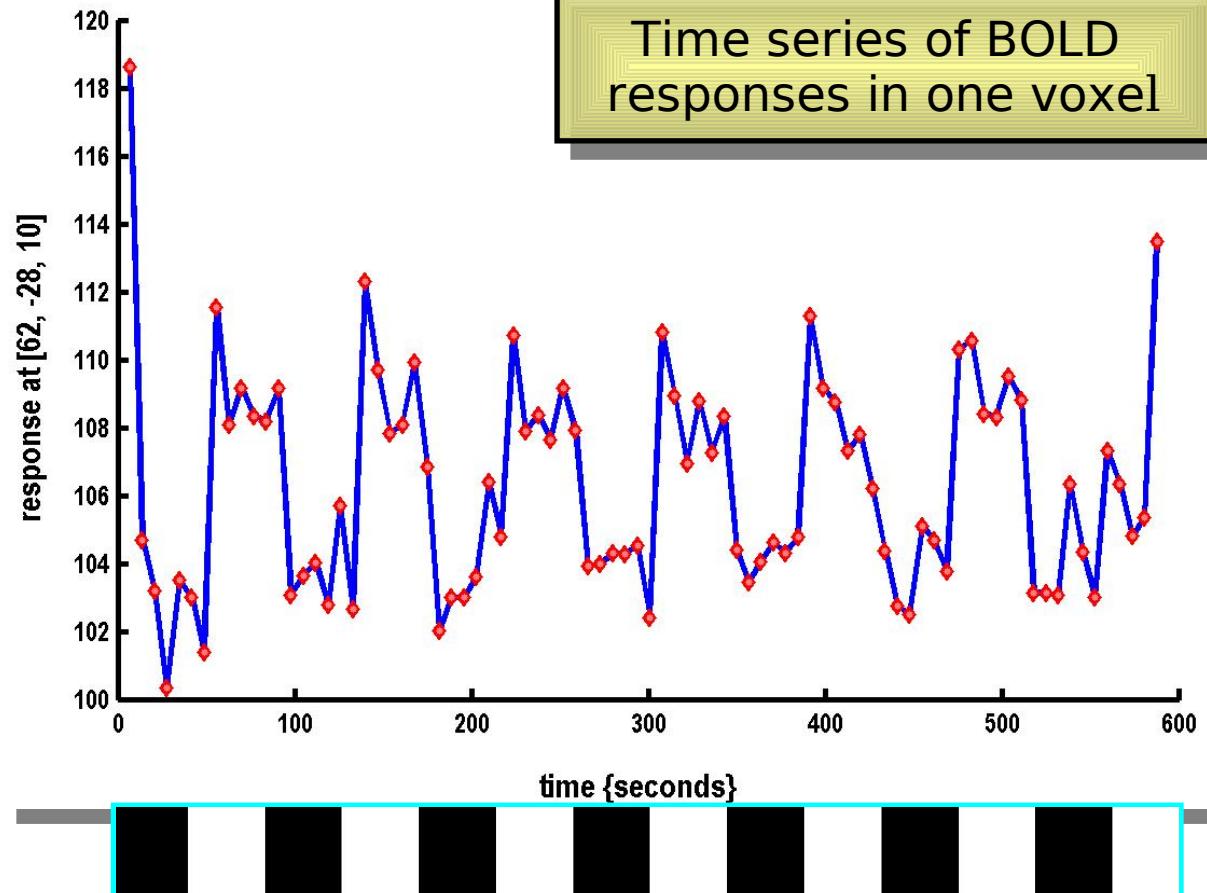
One session

Passive word listening
versus rest

7 cycles of
rest and listening

Each epoch 6 scans
with 7 sec TR

Time series of BOLD
responses in one voxel



Question: Is there a change in the BOLD
response between listening and rest?

Stimulus function

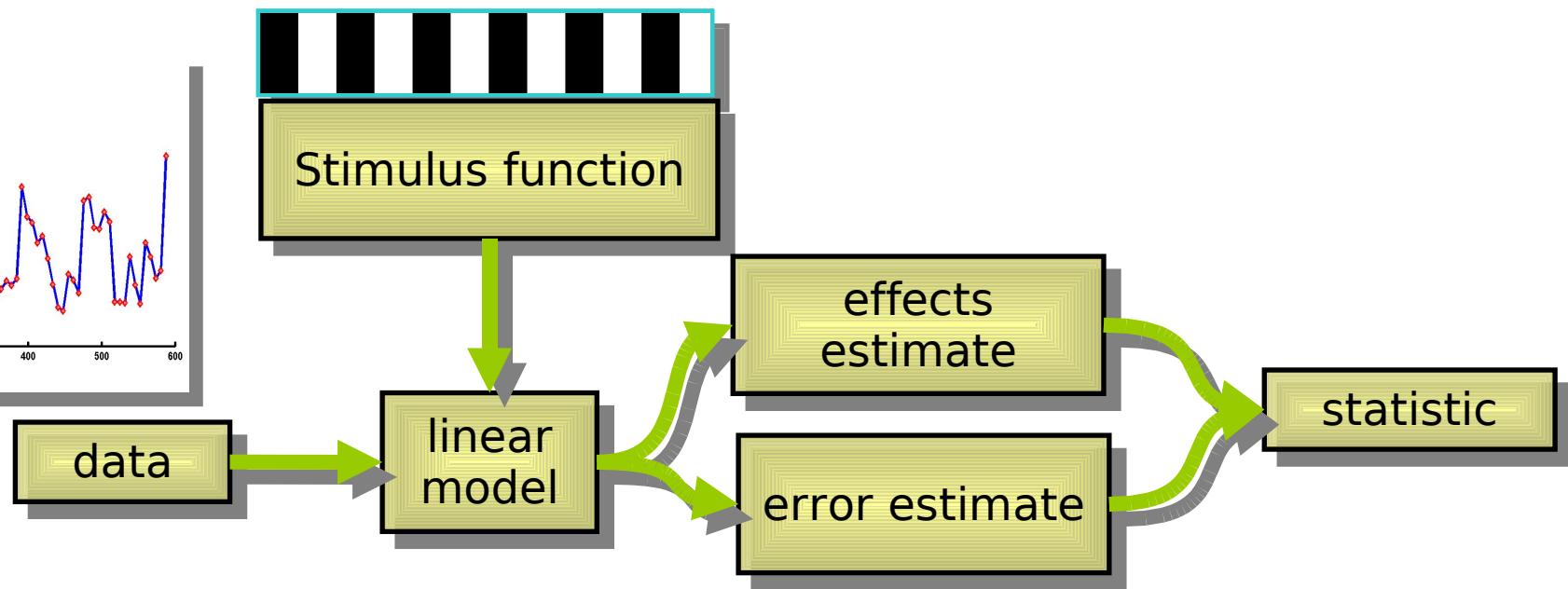
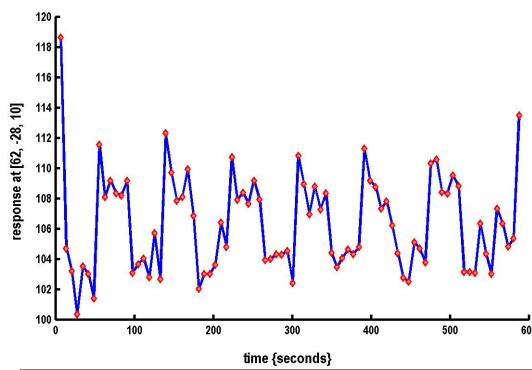
Why modelling?

Why?

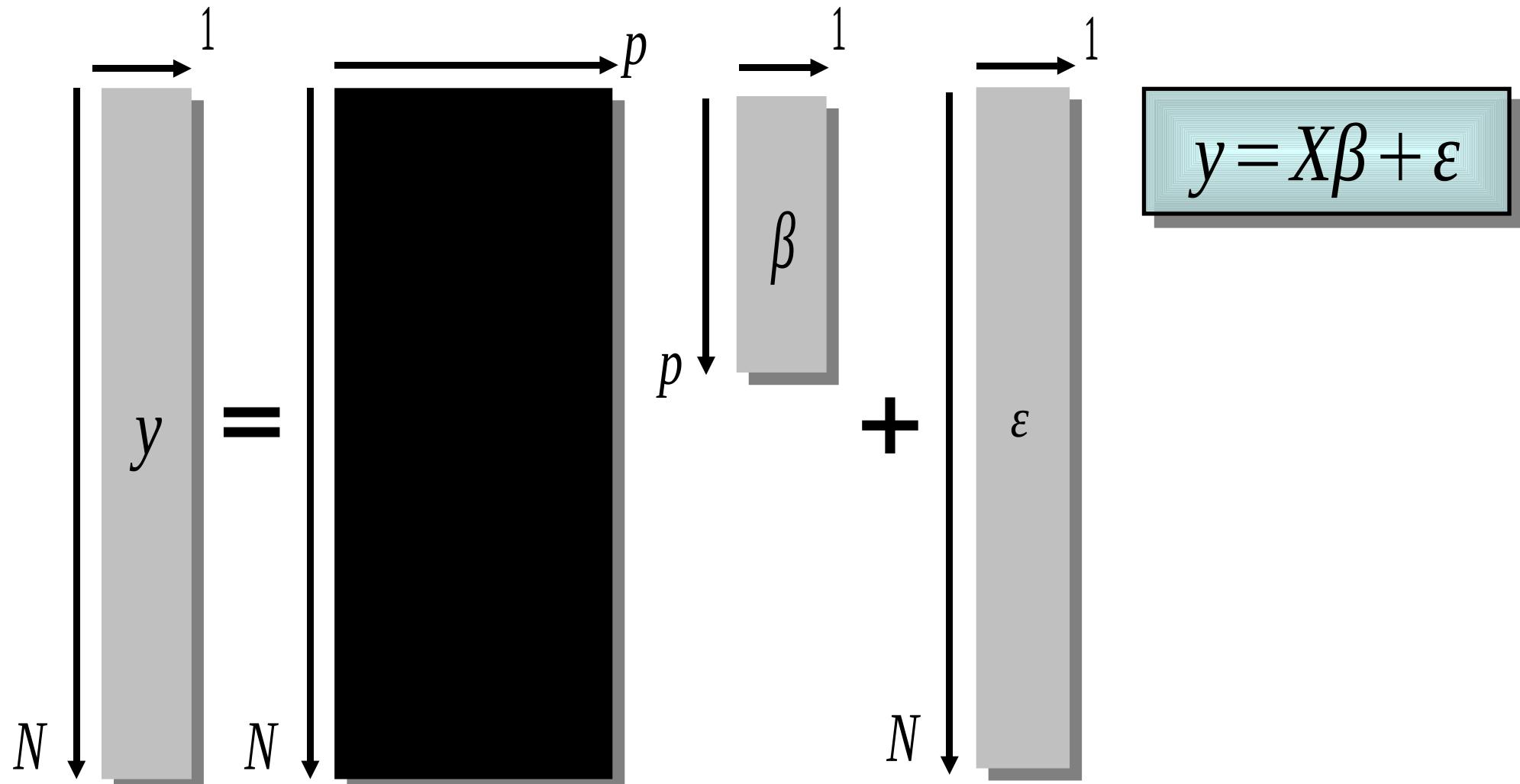
Make inferences about effects of interest

How?

1. Decompose data into effects and error
2. Form statistic using estimates of effects and error



General Linear Model



N : number of scans
 p : number of regressors

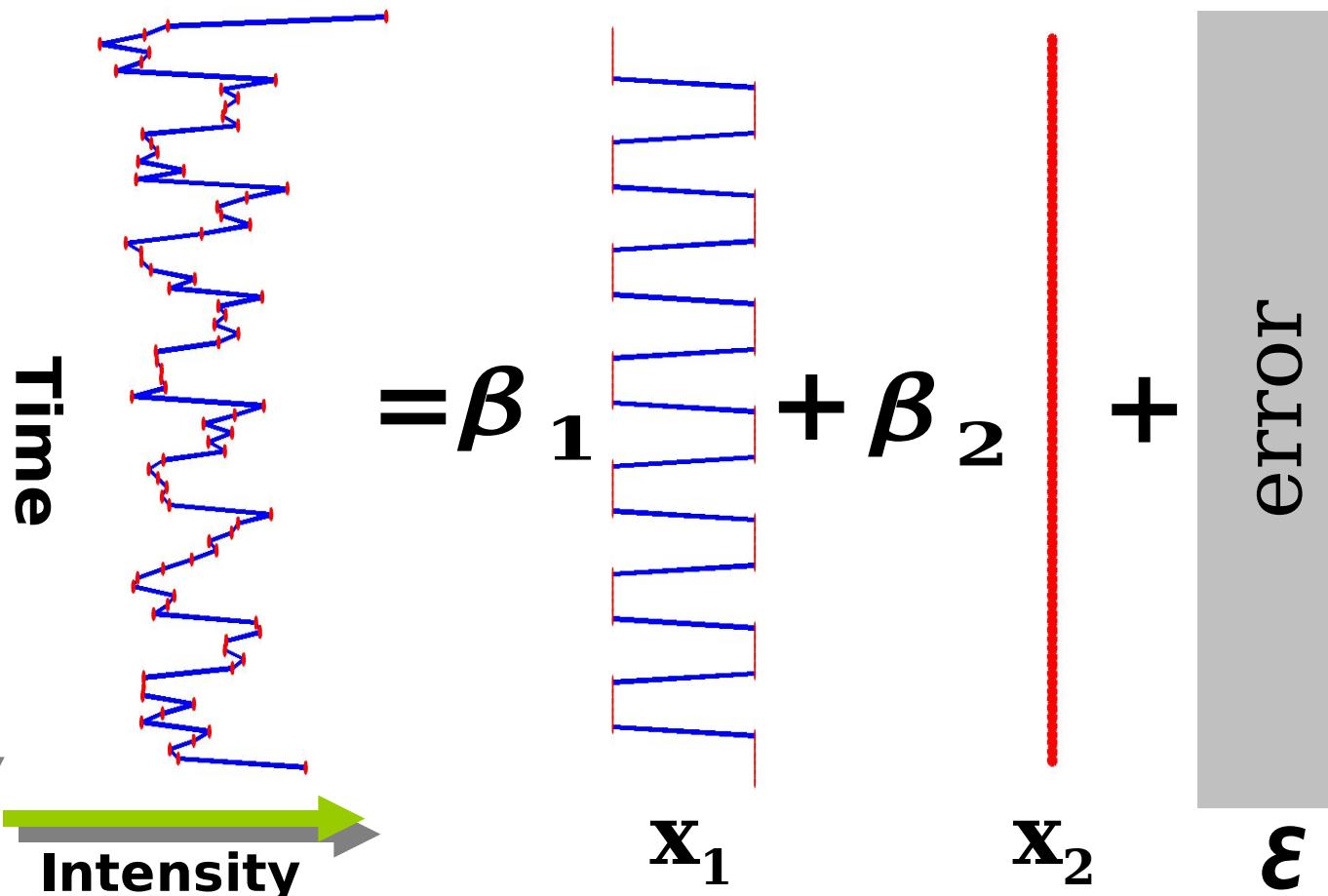
Model is specified by

- Design matrix X
- Assumptions about ε



General Linear Model

cea



$$\boldsymbol{\epsilon} \sim N(0, \sigma^2 I)$$

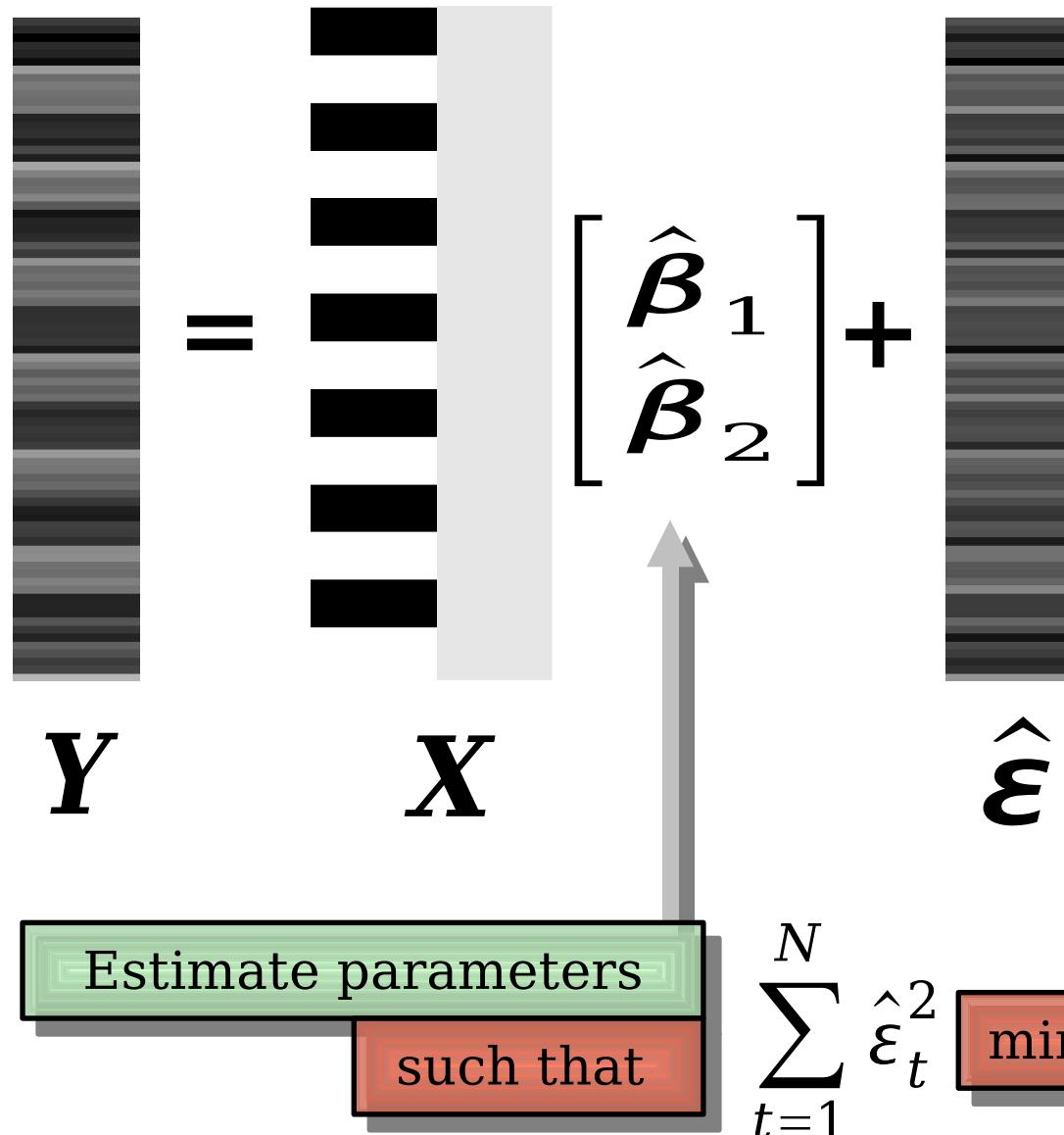
(error is normal
and iid)

Question: Is there a change in the BOLD response between listening and rest?

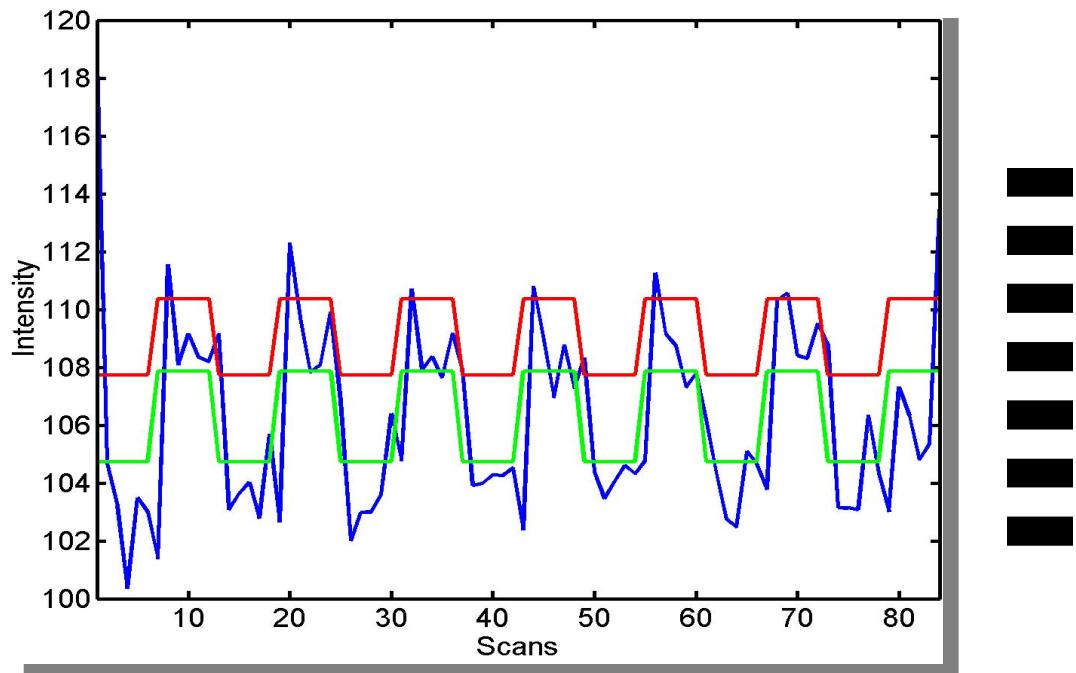
Hypothesis test: $\beta_1 = 0$?
(using t-statistic)



Parameter estimation



Estimation, example



Least squares estimate $\hat{\beta} = \begin{bmatrix} 3.15 \\ 104.74 \end{bmatrix}$

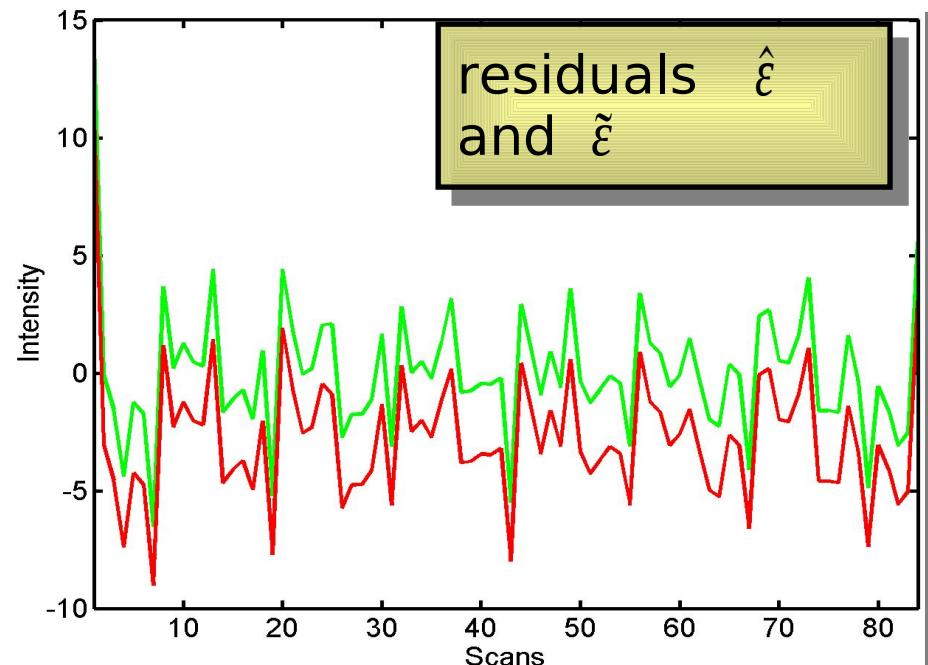
Another estimate $\tilde{\beta} = \begin{bmatrix} 2.65 \\ 107.74 \end{bmatrix}$



$$Y = X \beta + \varepsilon$$

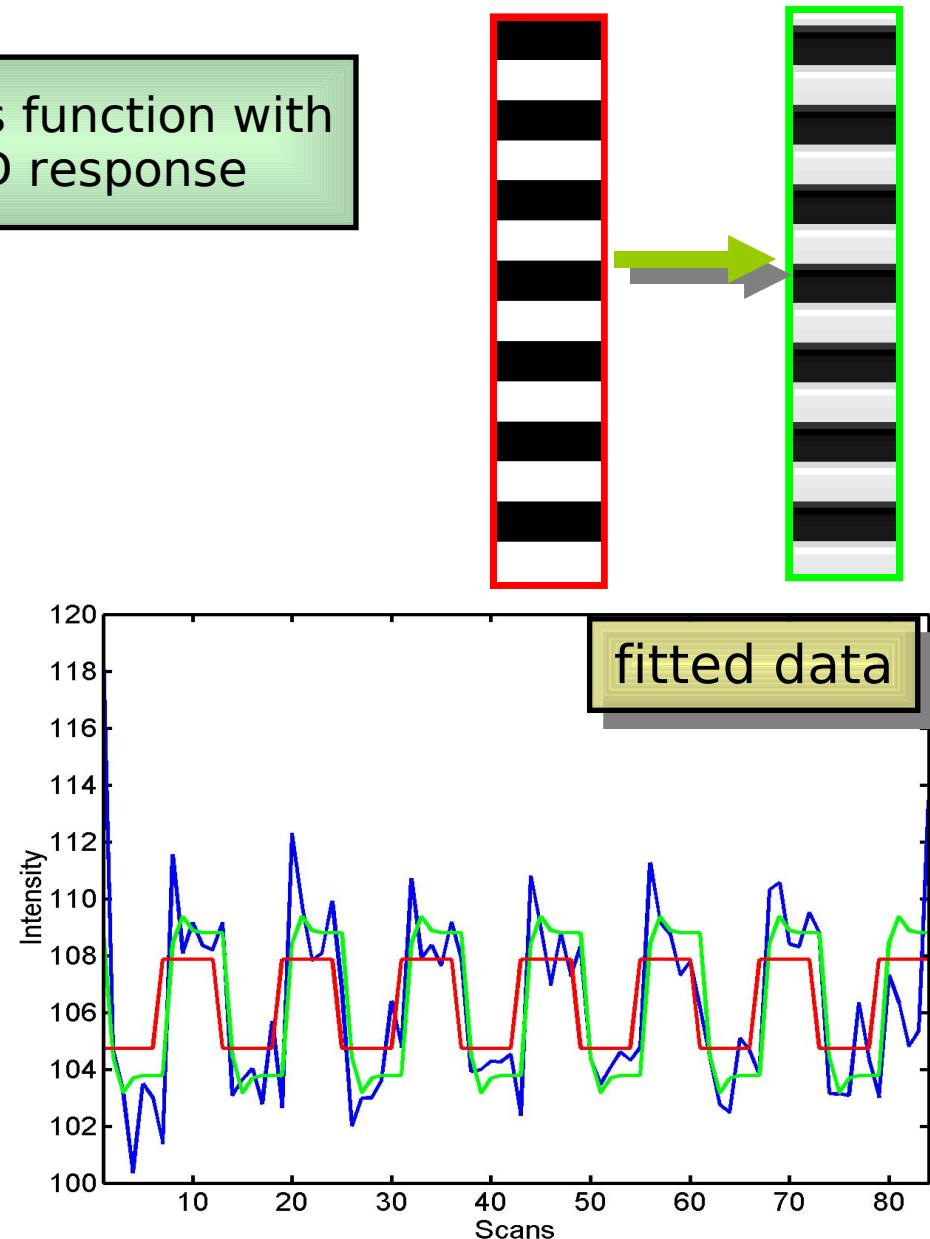
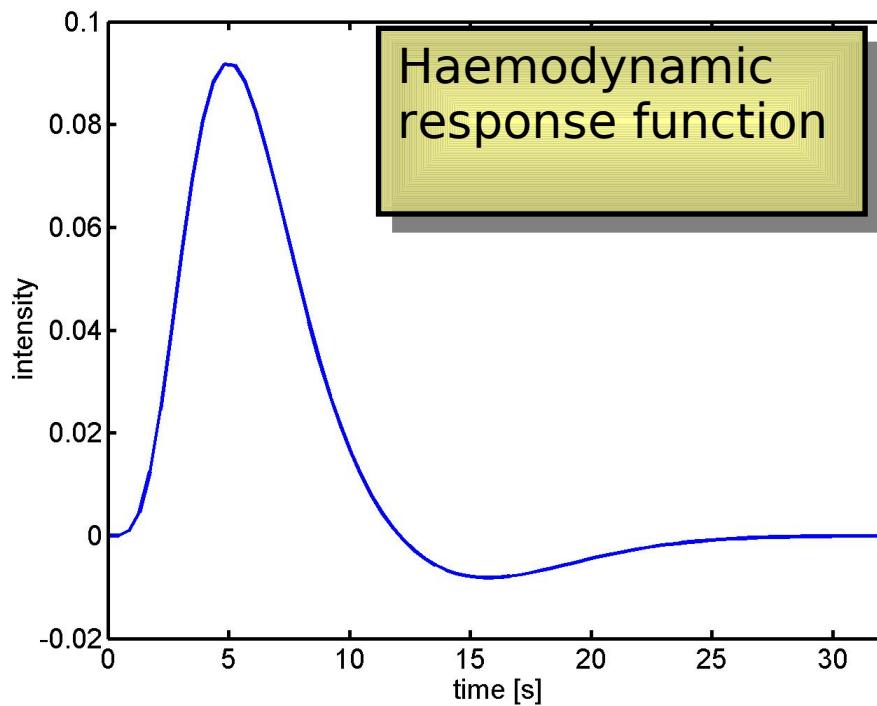
Assume iid error

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$



Improved model

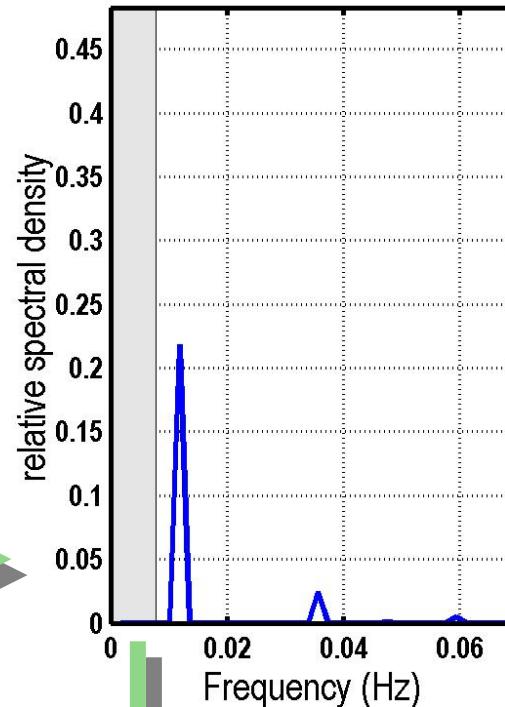
Convolve stimulus function with
model of BOLD response



High-pass filtering

$$Y = X \beta + \varepsilon$$

Frequency domain
128 second High-pass filter

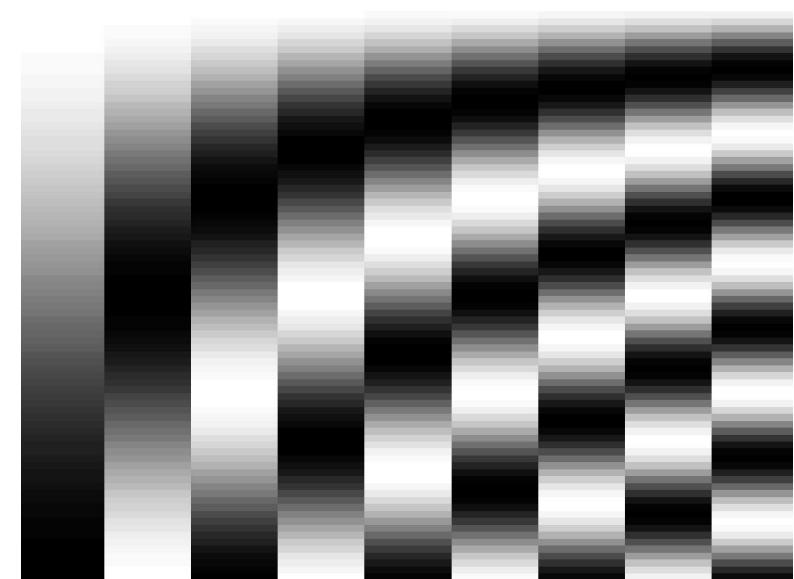


High-pass filter
implemented by
modelling low
frequencies

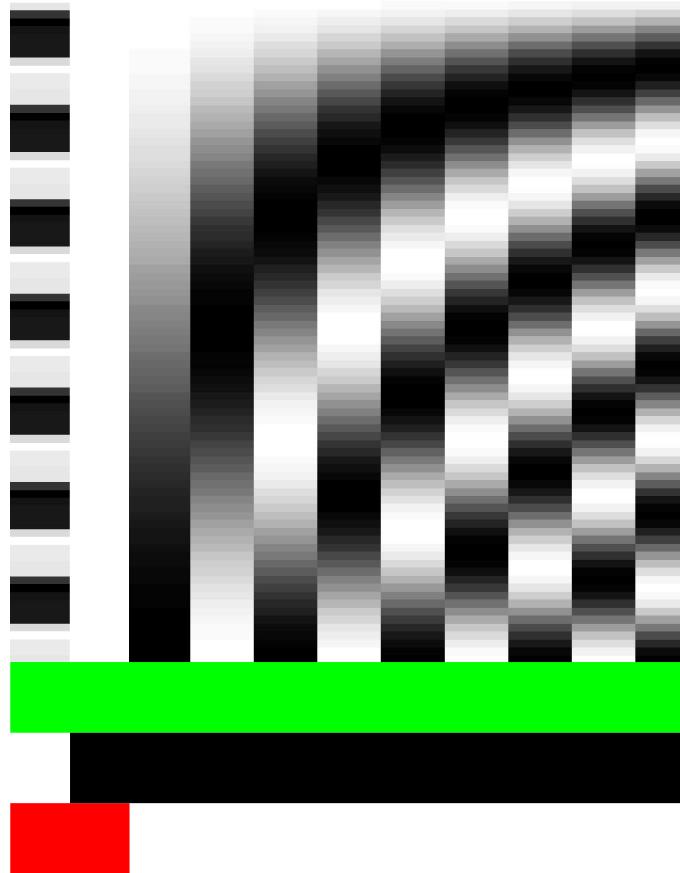
discrete cosine
transform set

N

p

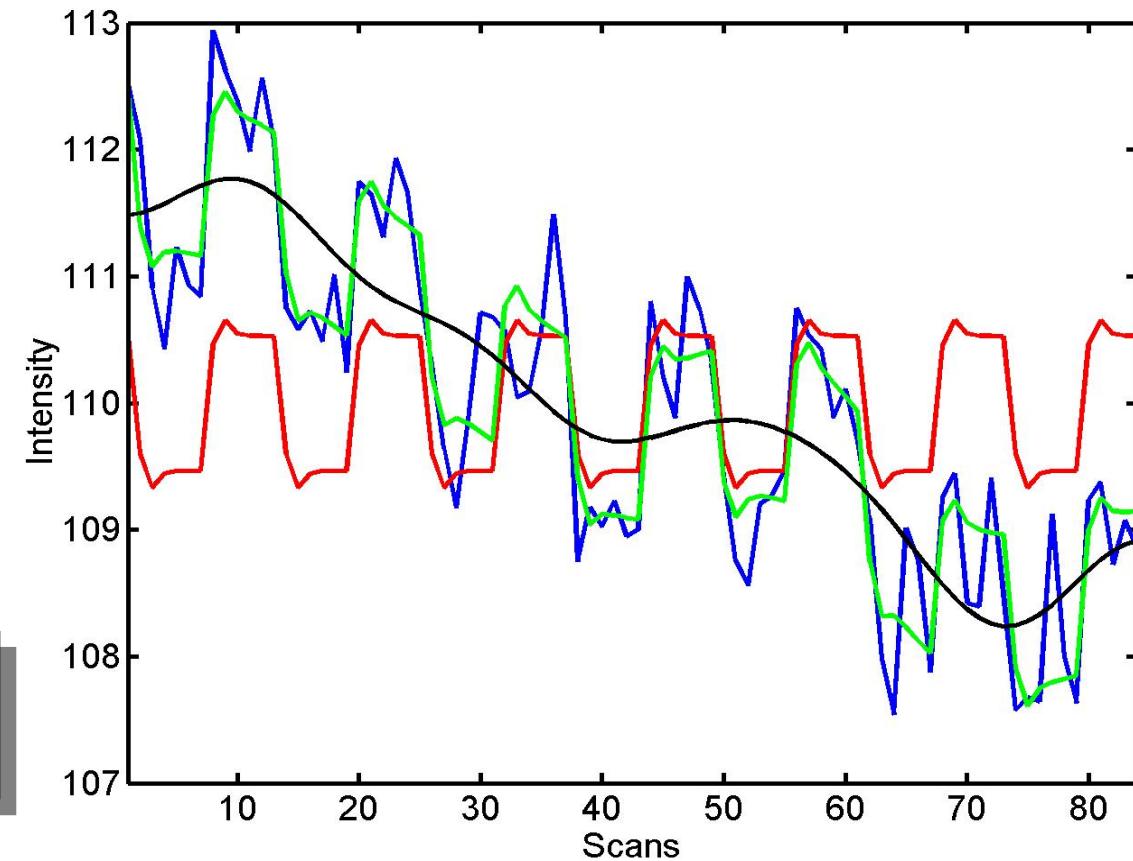


High-pass filtering



the data and three
different models

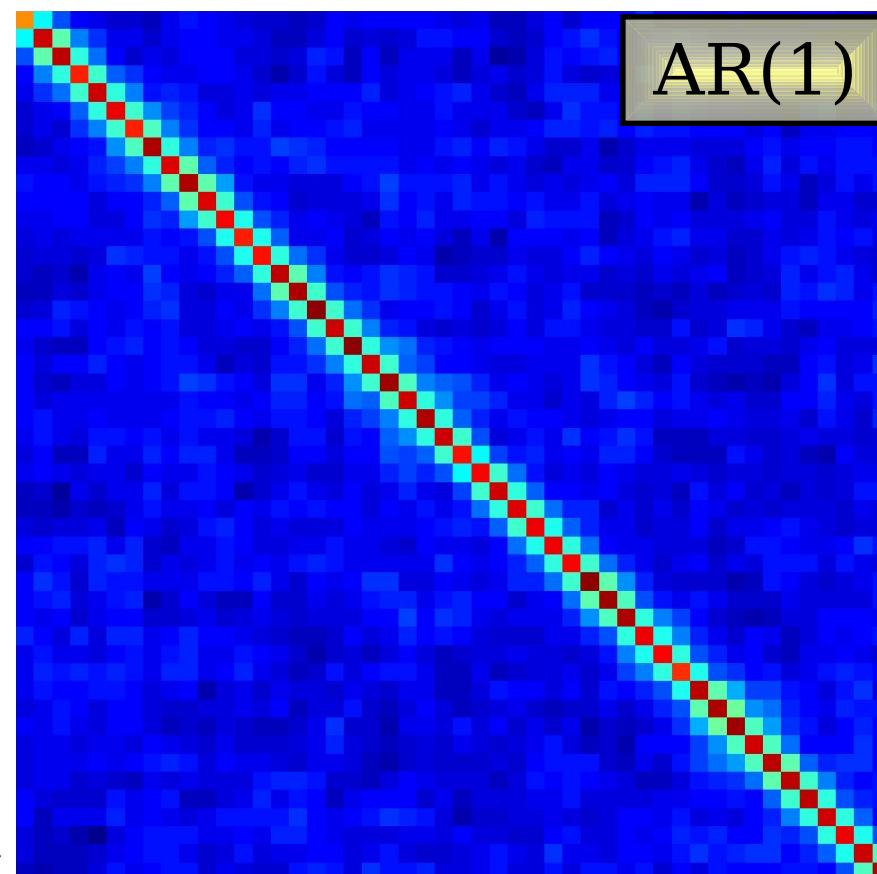
$$Y = X \beta + \varepsilon$$



Error covariance matrix

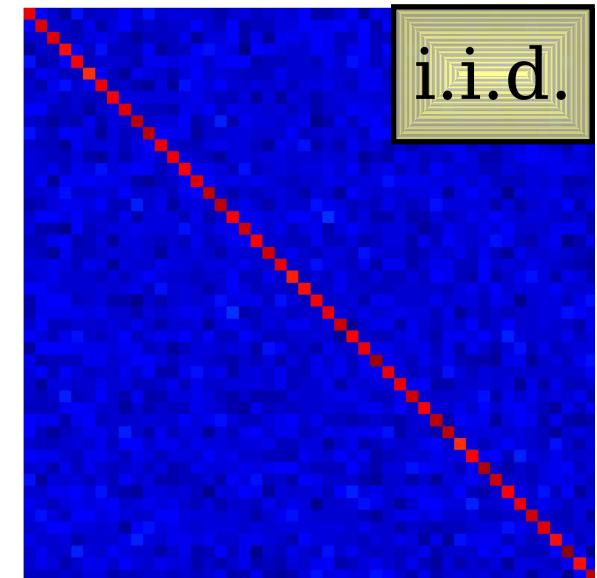
$$Y = X\beta + \varepsilon$$

$$\text{Cov}(\varepsilon)$$



N

$$V_{ij} = \sigma^2 \frac{a^{|i-j|}}{1-a^2}$$



sampled error covariance matrix

Serial correlations



Maximum likelihood solution

- Least Mahalanobis distance (Gaussian assumption)

- Assume V is known up to a scalar factor: $V = \sigma^2 W$
- The ML effect estimator minimizes the Mahalanobis distance

$$d_{maha}^2 = (Y - X\beta)^t W^{-1} (Y - X\beta) = \|W^{-\frac{1}{2}}(Y - X\beta)\|^2$$

- To see this, note that:

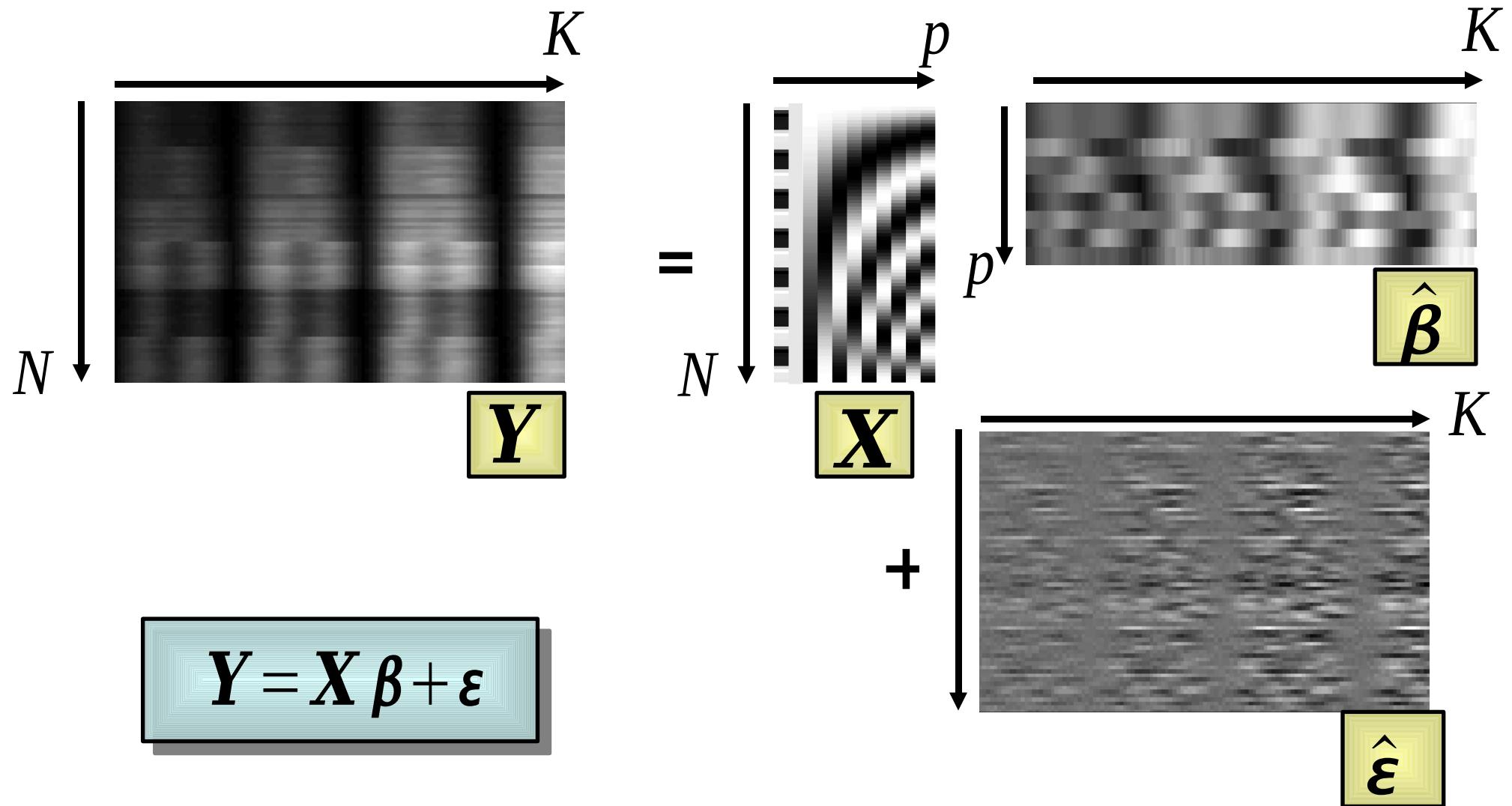
$$W^{-\frac{1}{2}} Y = W^{-\frac{1}{2}} X\beta + \sigma^2 N(0, I_n)$$

- Therefore, only need to pre-whiten the data and the design matrix

$$Y \leftarrow W^{-\frac{1}{2}} Y, \quad X \leftarrow W^{-\frac{1}{2}} X$$



Mass-univariate approach



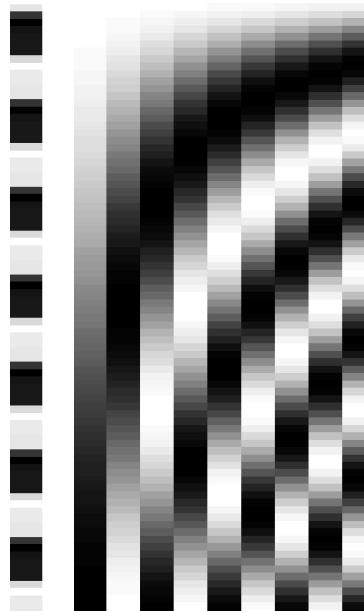
$$Y = X \hat{\beta} + \hat{\epsilon}$$



Inference - t-statistic

$$Y = X \beta + \epsilon$$

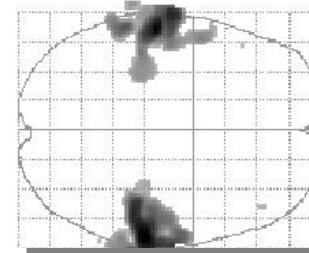
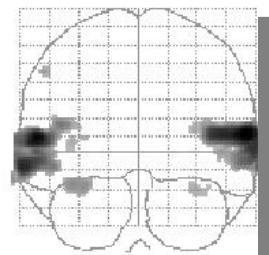
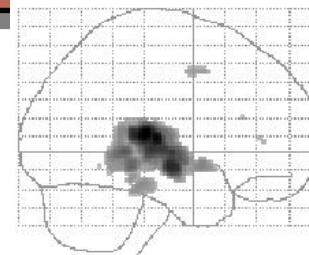
$$\begin{matrix} c = +1 & 0 & 0 & 0 & 0 & 0 & 0 \\ & 0 & 0 & 0 \end{matrix}$$



boxcar parameter > 0 ?

Null hypothesis: $\beta_1 = 0$

$$t = \frac{\mathbf{c}^T \hat{\boldsymbol{\beta}}}{\sqrt{\text{Var}(\mathbf{c}^T \hat{\boldsymbol{\beta}})}}$$



SPM{T₇₃}



t-statistic - Computations

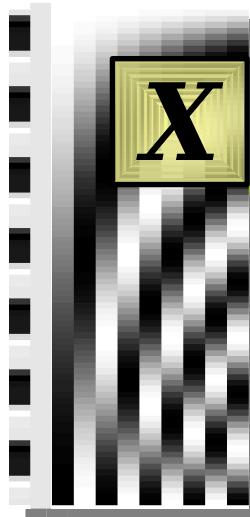
$$Y = X \beta + \varepsilon$$

$\hat{\beta}$ least squares estimates

$$c = +1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0$$

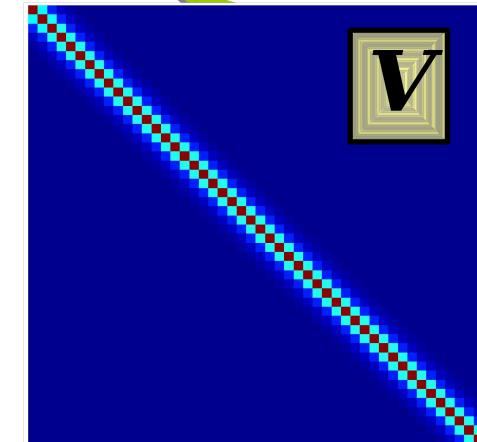
$$t = \frac{c^t \hat{\beta}}{\sqrt{Var(c^t \hat{\beta})}}$$

$$Var(c^t \hat{\beta}) = \hat{\sigma}^2 c^t X V X^{-t} c$$



$$\hat{\sigma}^2 = \frac{\sum_{t=1}^N (y_t - [X \hat{\beta}]_t)^2}{df}$$

compute df using Satterthwaite approximation



ReML



Hypothesis Testing

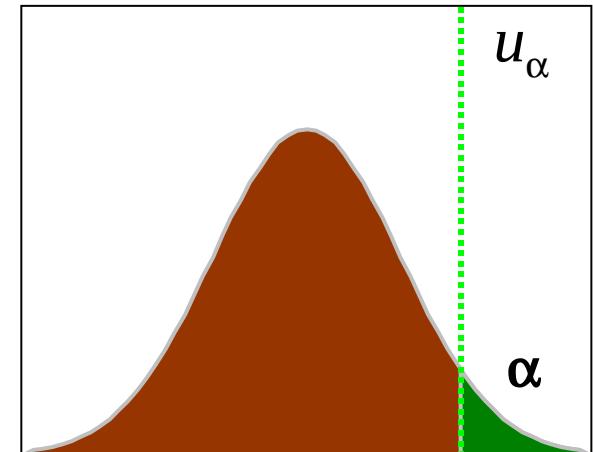
□ Type I Error α :

Acceptable *false positive rate* α .

Level \Rightarrow threshold u_α

Threshold u_α controls the false positive rate

$$\alpha = p(T > u_\alpha | H_0)$$



Null Distribution of T

□ P-value:

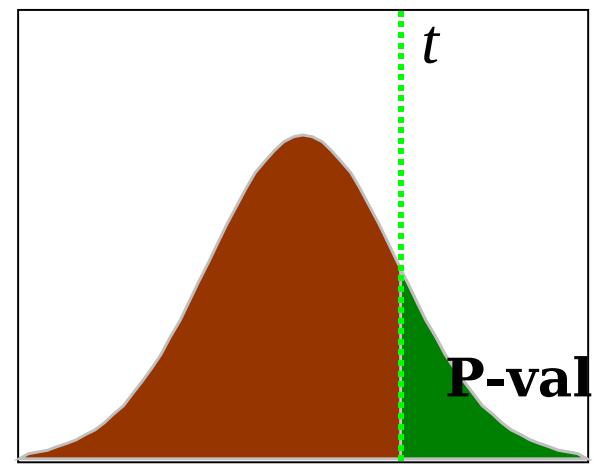
A *p*-value summarises evidence against H_0 .

This is the chance of observing value more extreme than t under the null hypothesis.

$$p(T > t | H_0)$$

□ The conclusion about the hypothesis:

We reject the null hypothesis in favour of the alternative hypothesis if $t > u_\alpha$

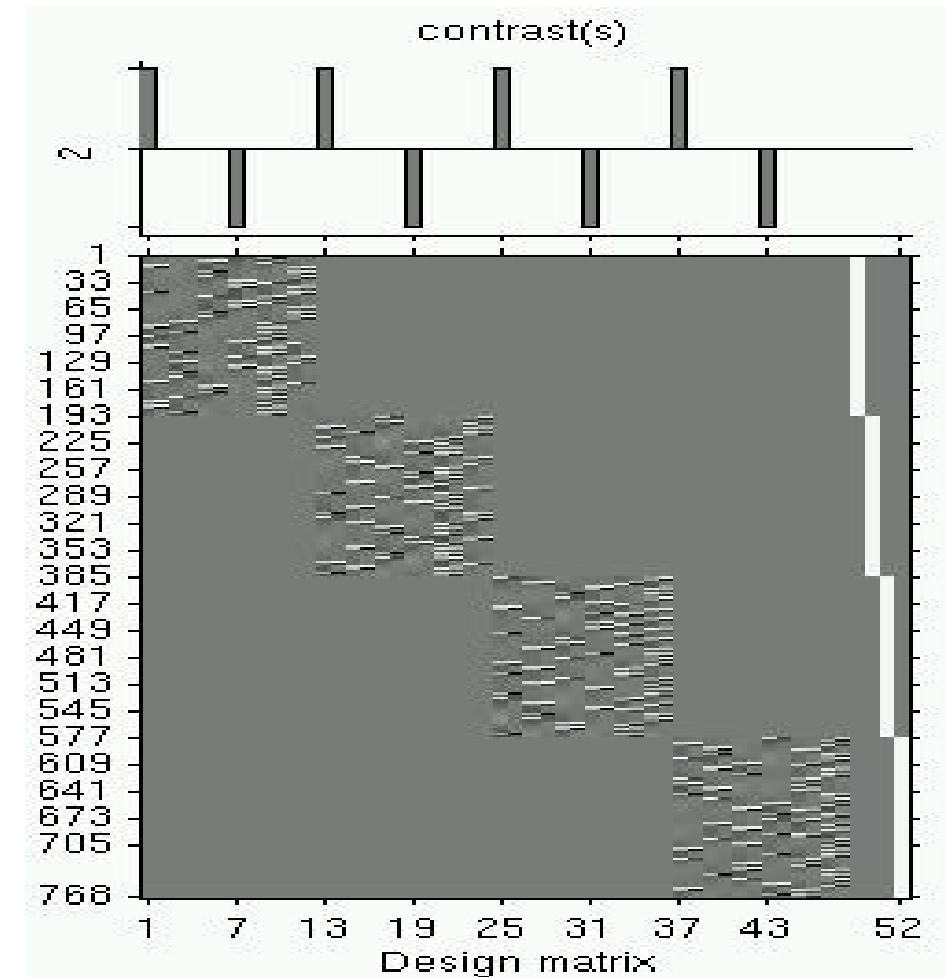
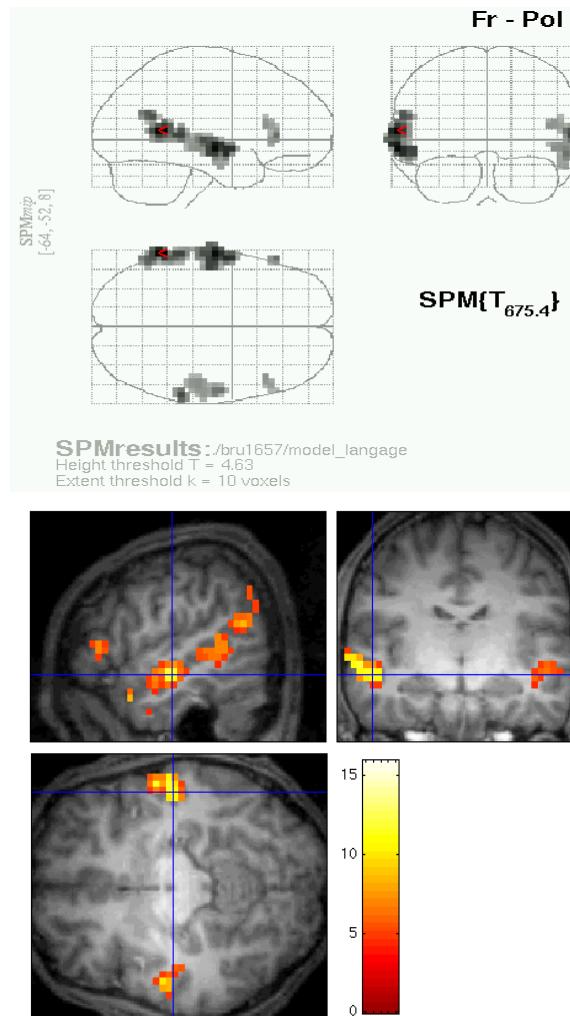


Null Distribution of T



Example of fMRI model

- A language comprehension study [Pallier et al, 2002]



Part I - Mapping brain activity

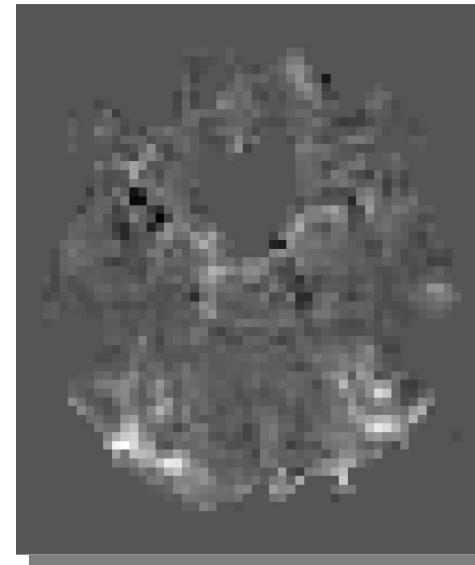
- A) A tour about the GLM framework
- B) What kind of regularization?
- C) Numerical Bayesian inference methods



Experimental evidence

Even without applied spatial smoothing, activation maps (and maps of eg. AR coefficients) have spatial structure.

Contrast



AR(1)



- ⇒ Definition of a spatial prior via Gaussian Markov Random Field
- ⇒ Automatic spatial regularisation of Regression coefficients and AR coefficients

Bayesian fMRI



Bayesian fMRI

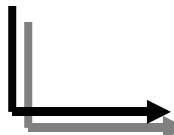
General Linear Model:

$$Y = X\beta + \varepsilon$$

with $\varepsilon \sim N(0, C_\varepsilon)$

What are the priors?

- In “*classical*” SPM, no (flat) priors
- In “*full*” Bayes, priors might be from theoretical arguments or from independent data
- In “*empirical*” Bayes, priors derive from the same data, assuming a hierarchical model for generation of the data



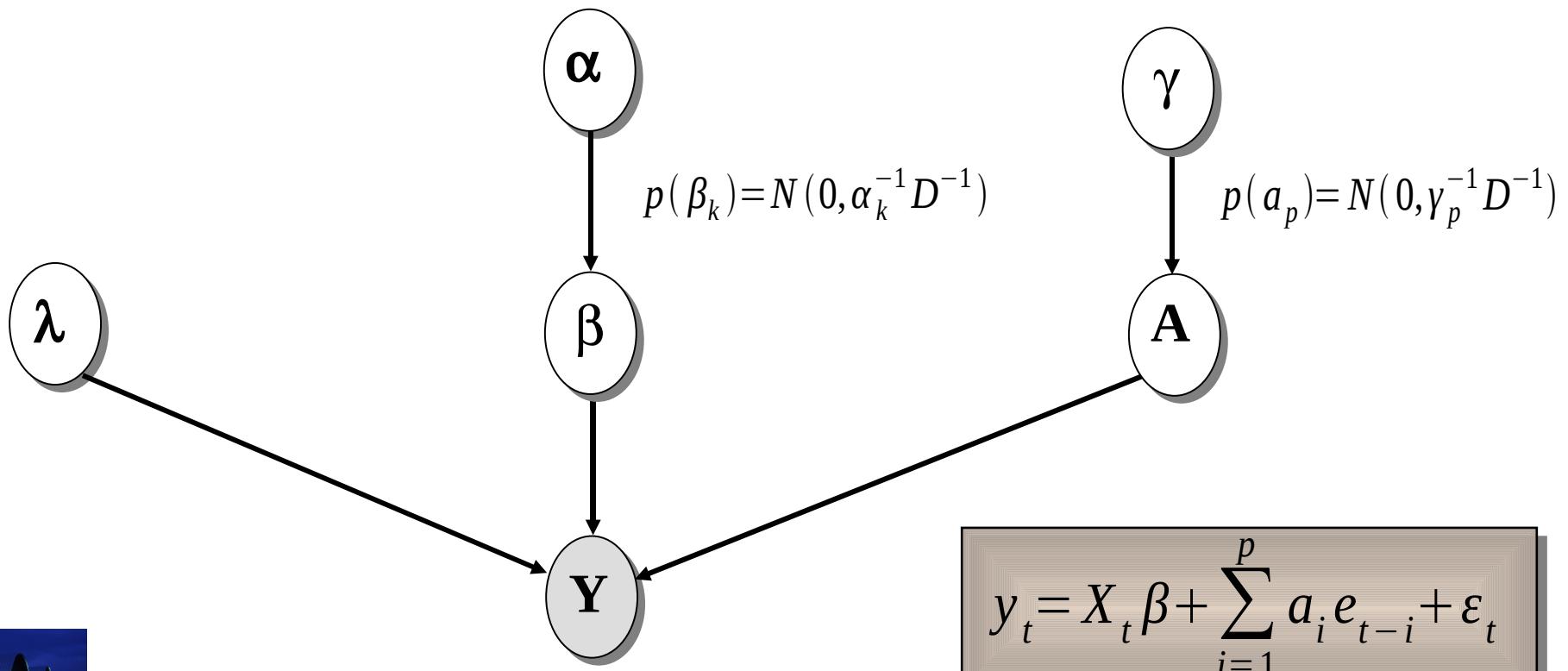
Parameters of one level can be made priors on distribution of parameters at lower level



The generative model

General Linear Model with Auto-Regressive error terms (GLM-AR):

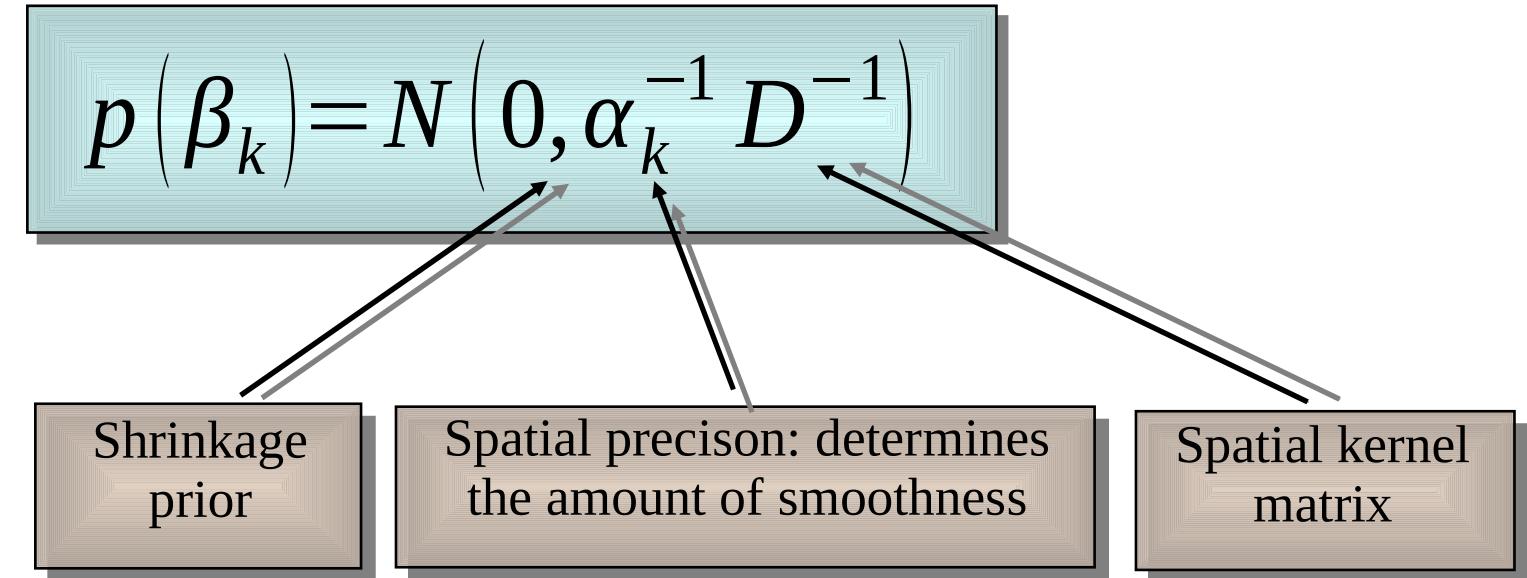
$$Y = X \beta + E \text{ where } E \text{ is an AR}(p)$$



Spatial prior

Over the regression coefficients:

[Penny et al, NeuroImage, 2003, 2005]



Gaussian Markov Random Field priors D

$$D = \begin{bmatrix} 1 & & & \\ & 1 & & \\ & & \ddots & d_{ij} \\ & & d_{ji} & 1 \\ & & & & 1 \end{bmatrix}$$

$\left\{ \begin{array}{l} 1 \text{ on diagonal elements } d_{ii} \\ d_{ij} > 0 \text{ if voxels } i \text{ and } j \text{ are neighbors.} \\ 0 \text{ elsewhere} \end{array} \right.$

Same prior on the AR coefficients.



Prior, Likelihood and posterior

The prior:

$$p(\beta, A, \lambda, \alpha, \gamma) = \left(\prod_k p(\beta_k | \alpha_k) p(\alpha_k | q_1, q_2) \right) \left(\prod_p p(a_p | \gamma_p) p(\gamma_p | r_1, r_2) \right) \\ \left(\prod_n p(\lambda_n | u_1, u_2) \right)$$

The likelihood:

$$p(Y | \beta, A, \lambda) = \prod_n p(y_n | \beta_n, a_n, \lambda_n)$$

The posterior?

$$p(\beta | Y) ?$$

The posterior over β doesn't factorise over k or n .

- ⇒ Exact inference requires sampling techniques
- ⇒ Variational approximation achievable at lower cost



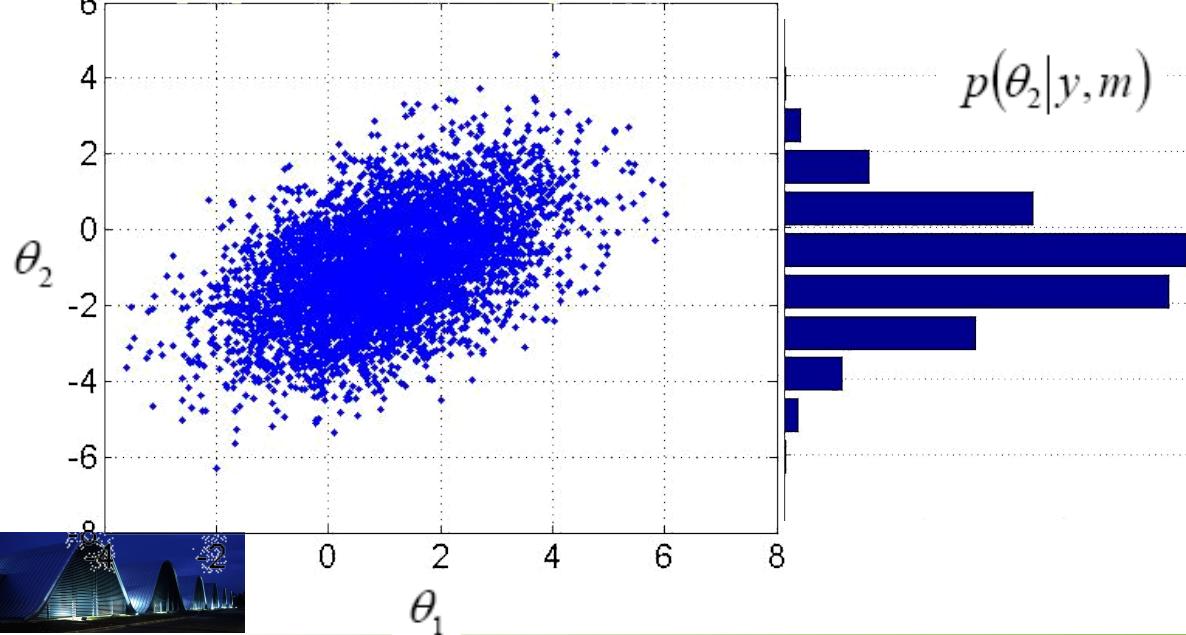
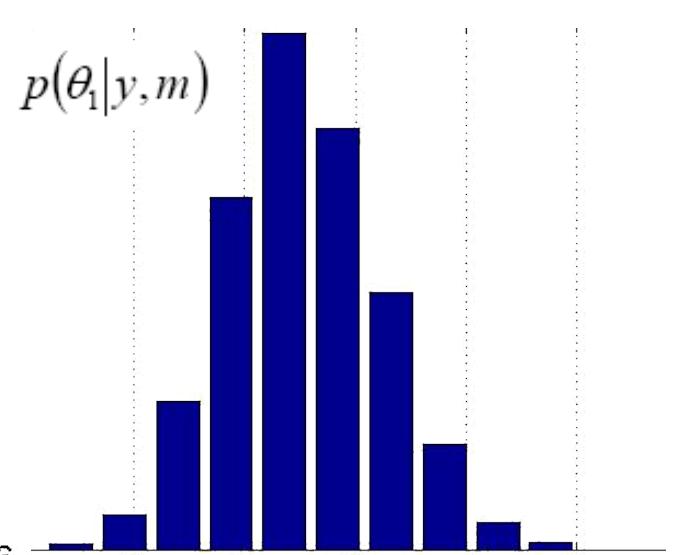
Part I - Mapping brain activity

- A) A tour about the GLM framework
- B) What kind of regularization?
- C) Numerical Bayesian inference methods

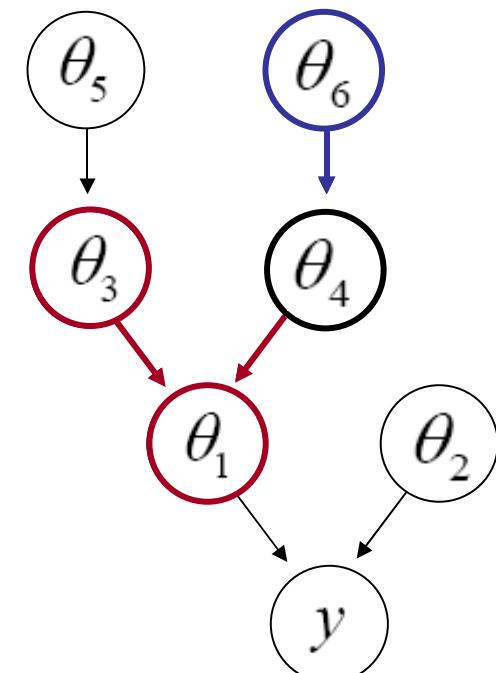


Sampling methods

MCMC example: Gibbs sampling



$$p(\theta_i|y, \theta_{j \neq i}) \propto \frac{p(\theta_i|par(\theta_i))}{\prod_{j=ch(i)} p(\theta_j|par(\theta_j))}$$



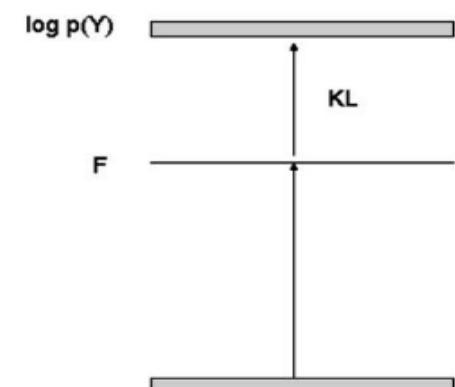
Variational Bayes

- Central quantity of Bayesian learning $p(\boldsymbol{\theta} | \mathbf{Y})$ with $\boldsymbol{\theta} = \{\beta, A, \lambda\}$
- Log-evidence of the model or integrated likelihood

$$\begin{aligned}
 \forall \text{ pdf } q, \log p(\mathbf{Y}) &= \int q(\boldsymbol{\theta} | \mathbf{Y}) \log p(\mathbf{Y}) d\boldsymbol{\theta} \\
 &= \int q(\boldsymbol{\theta} | \mathbf{Y}) \log \frac{p(\mathbf{Y}, \boldsymbol{\theta})}{p(\boldsymbol{\theta} | \mathbf{Y})} d\boldsymbol{\theta} \\
 &= \int q(\boldsymbol{\theta} | \mathbf{Y}) \log \left[\frac{q(\boldsymbol{\theta} | \mathbf{Y}) p(\mathbf{Y}, \boldsymbol{\theta})}{p(\boldsymbol{\theta} | \mathbf{Y}) q(\boldsymbol{\theta} | \mathbf{Y})} \right] d\boldsymbol{\theta} \\
 &= \int q(\boldsymbol{\theta} | \mathbf{Y}) \log \frac{p(\boldsymbol{\theta}, \mathbf{Y})}{q(\boldsymbol{\theta} | \mathbf{Y})} d\boldsymbol{\theta} + \int q(\boldsymbol{\theta} | \mathbf{Y}) \log \frac{q(\boldsymbol{\theta} | \mathbf{Y})}{p(\boldsymbol{\theta} | \mathbf{Y})} d\boldsymbol{\theta} \\
 &= \mathcal{F} + D(q || p_{\boldsymbol{\theta} | \mathbf{Y}})
 \end{aligned}$$

- \mathcal{F} is a lower bound of the model evidence

$$\log p(\mathbf{Y}) = \mathcal{F} \iff q(\boldsymbol{\theta} | \mathbf{Y}) = p(\boldsymbol{\theta} | \mathbf{Y})$$



Variational Bayes (cont'd)

- Aim of VB : maximize \mathcal{F}
 - Make the approximate posterior $q(\boldsymbol{\theta} | \mathbf{Y})$ as close as possible to the true posterior $p(\boldsymbol{\theta} | \mathbf{Y})$
- Practical efficient algorithm:
 - Ensure tractability of integrals in \mathcal{F}
 - Generic procedure: mean-field approximation

$$q(\boldsymbol{\theta} | \mathbf{Y}) = \prod q(\theta_i | \mathbf{Y})$$

θ_i = ith group of parameters i

- Maximizers of \mathcal{F} **[Lappalainen and Miskin, 2000]**

$$q(\theta_i | \mathbf{Y}) = \frac{\exp[I(\theta_i)]}{\int \exp[I(\theta_i)] d\theta_i}$$

$$\text{with } I(\theta_i) = \int q(\boldsymbol{\theta}^{\setminus i} | \mathbf{Y}) \log p(\mathbf{Y}, \boldsymbol{\theta}) d\boldsymbol{\theta}^{\setminus i}$$



Variational Bayes (exemple)

Approximate posteriors that allows for *factorisation*

$$q(\beta, A, \lambda, \alpha, \gamma) = \left(\prod_k q(\alpha_k | Y) \right) \left(\prod_p q(\gamma_p | Y) \right) \left(\prod_n q(\beta_n | Y) q(a_n | Y) q(\lambda_n | Y) \right)$$

Variational Bayes Algorithm

```

Initialisation
While ( $\Delta F > \text{tol}$ )
    Update Suff. Stats. for  $\beta$ 
    Update Suff. Stats. for  $A$ 
    Update Suff. Stats. for  $\lambda$ 
    Update Suff. Stats. for  $\alpha$ 
    Update Suff. Stats. for  $\gamma$ 
End

```



Variational Bayes (exemple)

Approximate posteriors that allows for *factorisation*

$$q(\beta, A, \lambda, \alpha, \gamma) = \left(\prod_k q(\alpha_k | Y) \right) \left(\prod_p q(\gamma_p | Y) \right) \left(\prod_n q(\beta_n | Y) q(a_n | Y) q(\lambda_n | Y) \right)$$

Regression coefficients

$$q(\mathbf{w}_n) = N(\mathbf{w}_n; \hat{\mathbf{w}}_n, \hat{\Sigma}_n)$$

$$\hat{\mathbf{w}}_n = \hat{\Sigma}_n (\bar{\lambda}_n \tilde{\mathbf{b}}_n^T + \mathbf{r}_n)$$

$$\hat{\Sigma}_n = (\bar{\lambda}_n \tilde{\mathbf{A}}_n + \mathbf{B}_{nn})^{-1}$$

$$\mathbf{B} = \mathbf{H} (diag(\bar{\alpha}) \otimes \mathbf{S}^T \mathbf{S}) \mathbf{H}^T$$

$$\mathbf{r}_n = - \sum_{i=1, i \neq n}^N \mathbf{B}_{ni} \hat{\mathbf{w}}_i$$

AR coefficients

$$q(\mathbf{a}_n) = N(\mathbf{a}_n; \mathbf{m}_n, \mathbf{V}_n)$$

$$\mathbf{V}_n = (\bar{\lambda}_n \tilde{\mathbf{C}}_n + \beta \mathbf{I}_p)^{-1}$$

$$\mathbf{m}_n = \bar{\lambda}_n \tilde{\mathbf{D}}_n \mathbf{V}_n$$

Spatial precisions

$$q(\alpha) = \prod_{k=1}^K q(\alpha_k)$$

$$q(\alpha_k) = Ga(\alpha_k; g_k, h_k)$$

$$\frac{1}{g_k} = \frac{1}{2} \left[Tr(\hat{\Sigma}_k \mathbf{S}^T \mathbf{S}) + \hat{\mathbf{w}}_k^T \mathbf{S}^T \mathbf{S} \hat{\mathbf{w}}_k \right] + \frac{1}{q_1}$$

$$h_k = \frac{N}{2} + q_2$$

$$\bar{\alpha}_k = g_k h_k$$

Observation noise

$$q(\lambda_n) = Ga(\lambda_n; b_n, c_n)$$

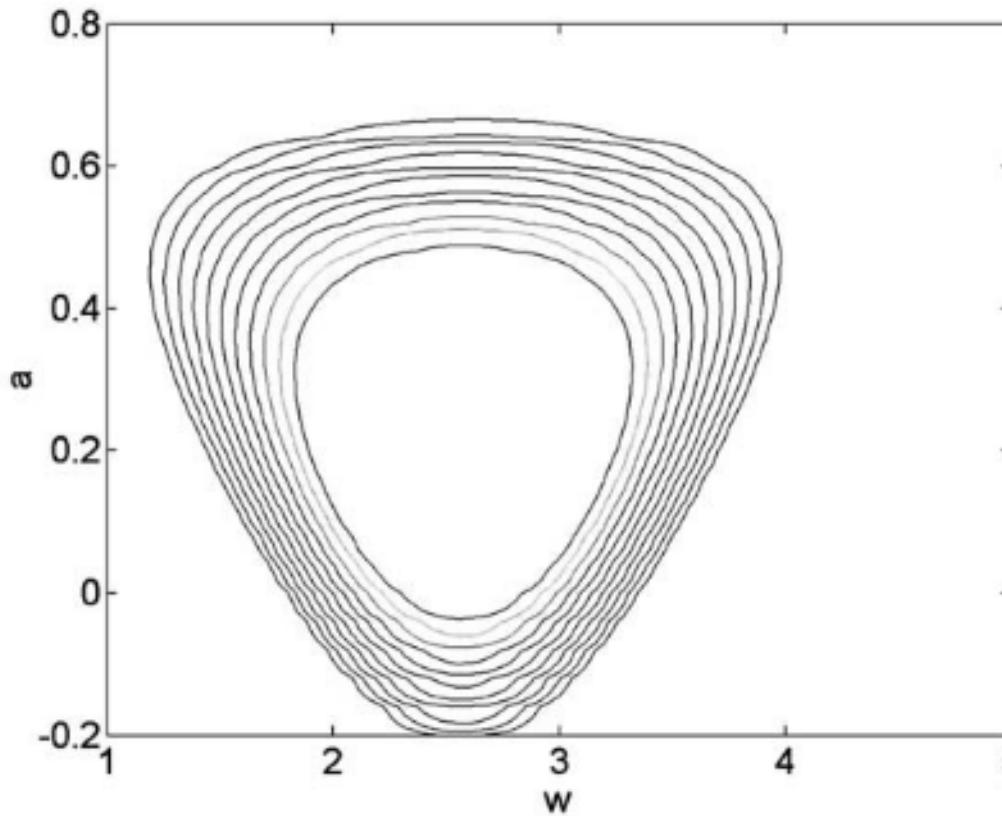
$$\frac{1}{b_n} = \frac{\tilde{G}_n}{2} + \frac{1}{u_1}$$

$$c_n = \frac{T}{2} + u_2$$

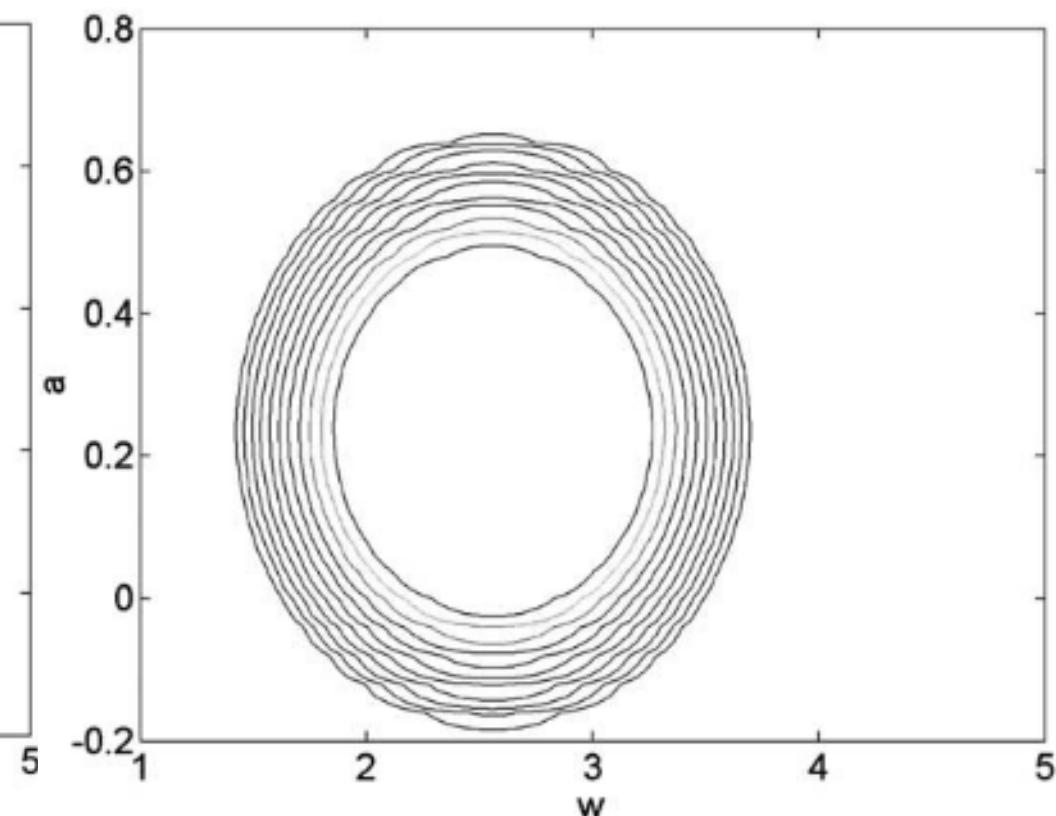


VB: approximation effect

iso-levels of the posterior $p(\theta | Y)$



iso-levels of the posterior $q(\theta | Y)$

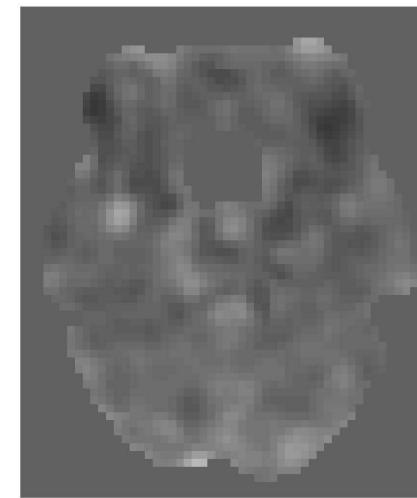


Accurate mode approximation
Error hidden in the **higher order moments**

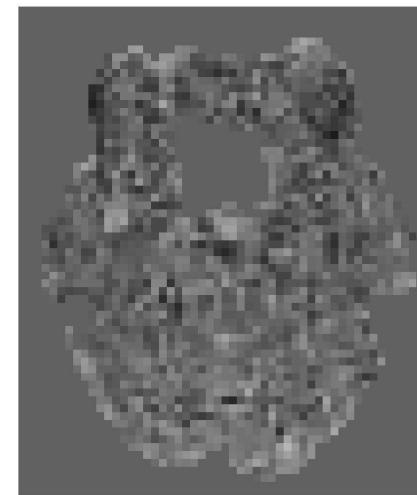


Event-related fMRI

Familiar vs.
unfamiliar faces



Smoothing



Global prior



Spatial Prior

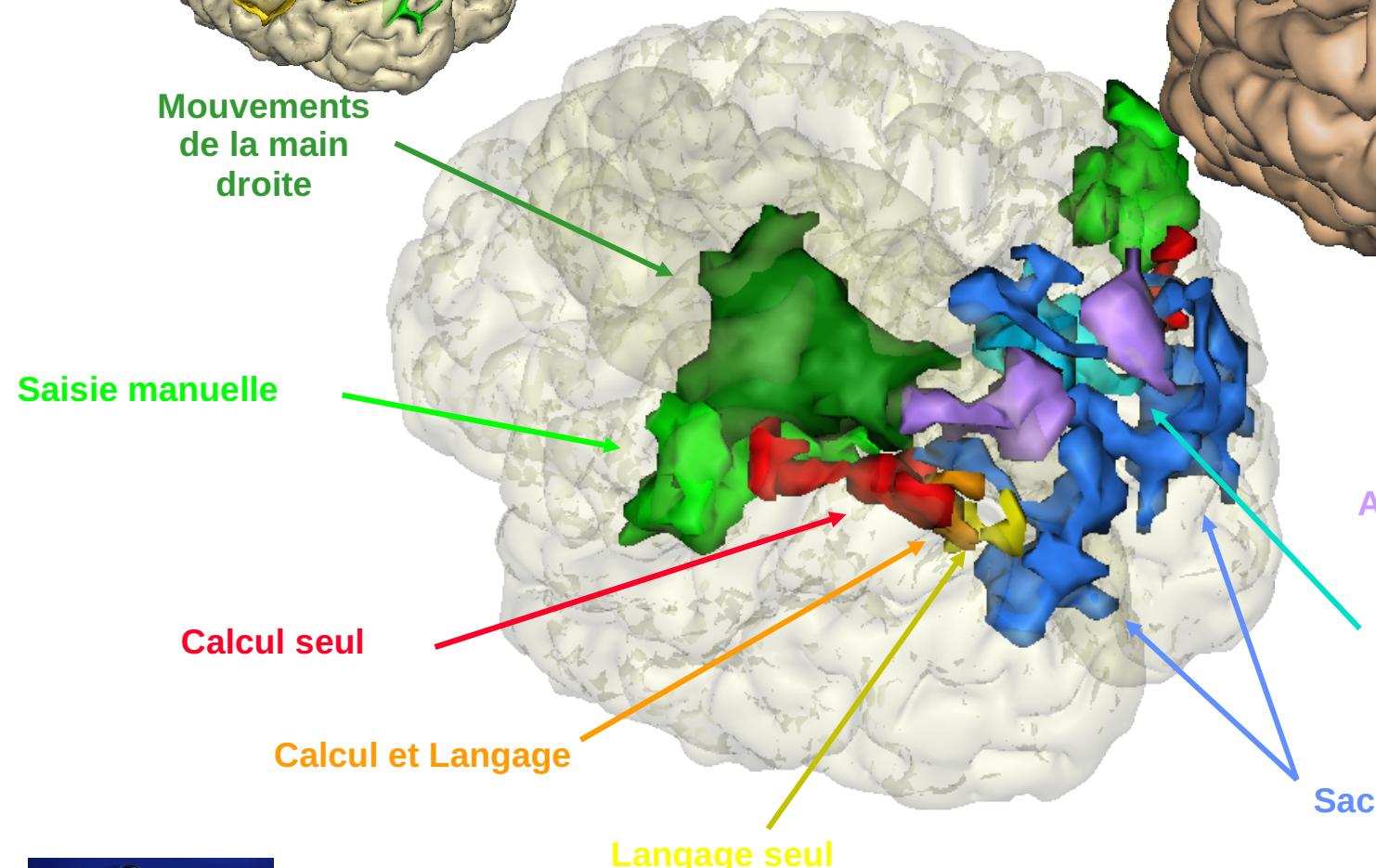
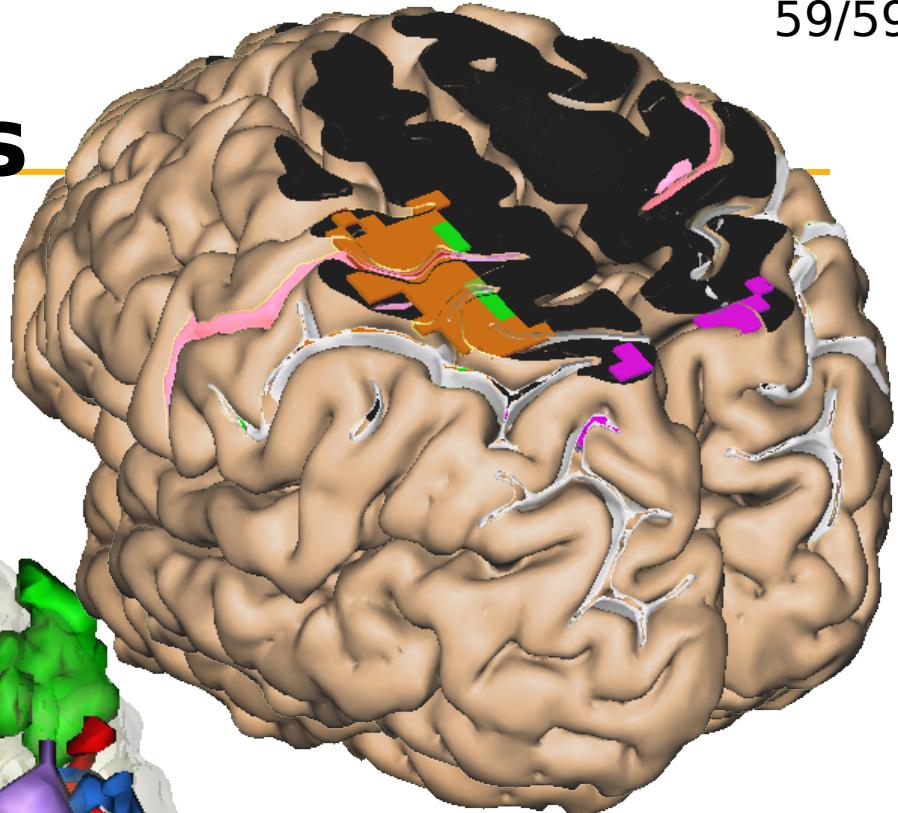
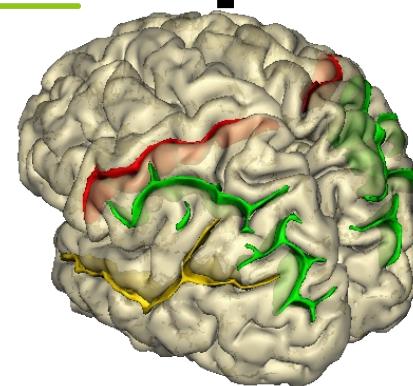


Summary

- Activation detection in fMRI
 - Preprocessings
 - A whole brain model of the BOLD signal
 - Statistical tests
- Bayesian gain
 - Don't smooth the data
 - Prefer spatial regularization



Mapping of the parietal circuits



O. Simon, D. Rivière (2000)